# Asymptotic normality of pattern counts in conjugacy classes 

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## Definitions

Permutations
Conjugacy invariant permutations
Patterns
Results
Uniform case: (Hofer)
Partial results: (Féray), (Hamaker and Rhoades) and
(Kammoun)
General case: (Dubach) and (Féray and Kammoun)
Proofs
Comparison techniques
Weighted dependency graphs
Universality (Aléa days)
I.I.D.

Random matrices
Longest increasing (decreasing) subsequence Conjugacy invariant permutations

## Permutation



Word:
21016987453
Descents
Peaks
Patterns
Longest increasing subsequence RSK


Cycles:
$(1,2,10,3)(4,6,8)(5,9)(7)$
Total number of cycles
Number of cycles of
length $i$
Conjugacy class

Matrix:

$$
\left[\begin{array}{llllllllll}
0 & 1 & 0 & 0 & 0 & 0 & 0 & 0 & 0 & 0 \\
0 & 0 & 0 & 0 & 0 & 0 & 0 & 0 & 0 & 1 \\
1 & 0 & 0 & 0 & 0 & 0 & 0 & 0 & 0 & 0 \\
0 & 0 & 0 & 0 & 0 & 1 & 0 & 0 & 0 & 0 \\
0 & 0 & 0 & 0 & 0 & 0 & 0 & 0 & 1 & 0 \\
0 & 0 & 0 & 0 & 0 & 0 & 0 & 1 & 0 & 0 \\
0 & 0 & 0 & 0 & 0 & 0 & 1 & 0 & 0 & 0 \\
0 & 0 & 0 & 1 & 0 & 0 & 0 & 0 & 0 & 0 \\
0 & 0 & 0 & 0 & 1 & 0 & 0 & 0 & 0 & 0 \\
0 & 0 & 1 & 0 & 0 & 0 & 0 & 0 & 0 & 0
\end{array}\right]
$$

Question: we fix the value of a function, we study another. Example in LIPN: Bassino et al.

- Condition: Separable i.e. 0 occurrence of the patterns 2413 and 3142
- Function to study: Longest increasing subsequence / proportion of other patterns.


## Cycle Structure and Spectrum

- \# total number of cycles
- $\#_{i}$ number of cycles of length $i$

If $0 \leq p<q$ and $\operatorname{GCD}(p, q)=1$, then

$$
\text { Multiplicity of eigenvalue } e^{\frac{p}{q} 2 \pi \mathrm{i}} \text { is } \sum_{r \geq 1} \#_{r q}(\sigma)
$$

In particular:

$$
\begin{gathered}
\#(\sigma)=\text { Multiplicity of eigenvalue } 1 \\
\operatorname{Tr}\left(\sigma^{k}\right)=\sum_{i \mid k} i \#_{i}(\sigma) \quad \text { and } \quad k \#_{k}(\sigma)=\sum_{i \mid k} \operatorname{Tr}\left(\sigma^{i}\right) \mu(i)
\end{gathered}
$$

Where $\mu(i)$ is the Möbius function defined as:

$$
\mu(i)= \begin{cases}0 & \text { if } i \text { is divisible by the square of a prime number, } \\ (-1)^{r} & \text { if } i \text { is the product of } r \text { distinct prime numbers }\end{cases}
$$

## Conjugacy Classes

The conjugacy class of $\sigma$ is $\left\{\pi \sigma \pi^{-1}, \pi \in \mathfrak{S}_{n}\right\}$.

## Theorem

Let $\sigma, \rho$ be two permutations.
There is equivalence between:

- $\sigma$ and $\rho$ are in the same conjugacy class
- $\sigma$ and $\rho$ have the same cycle structure, i.e., $\forall i \geq 1, \#_{i}(\sigma)=\#_{i}(\rho)$.
- $\sigma$ and $\rho$ have the same spectrum (considering multiplicities)
- $\forall i \geq 1, \operatorname{Tr}\left(\sigma^{i}\right)=\operatorname{Tr}\left(\rho^{i}\right)$.


## Conjugacy invariant

- Definition: $\sigma_{n}$ is conjugacy invariant if for all $\rho$,

$$
\rho \sigma_{n} \rho^{-1} \stackrel{d}{=} \sigma_{n}
$$

- $\sigma_{n}$ is conjugacy invariant if and only if $\mathbb{P}\left(\sigma_{n}=\sigma\right)$ is a function of the cycle structure of $\sigma$.


## Conjugacy invariant

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- $\sigma_{n}$ is conjugacy invariant if and only if $\mathbb{P}\left(\sigma_{n}=\sigma\right)$ is a function of the cycle structure of $\sigma$.
- Example 1: Ewens

$$
\mathbb{P}\left(\sigma_{n}=\sigma\right)=\frac{\theta^{\# \sigma}}{C_{n, \theta}}
$$

- Example 2: Uniform permutation within a conjugacy class.
- Example 3: Uniform Involutions / Derangements.

Morally: Conditioned on the cycle structure, the permutation is chosen uniformly.

## Descents

We denote by $D(\sigma)=\{i: \sigma(i+1)<\sigma(i)\}$.
We assume that $\left(\sigma_{n}\right)_{n \geq 1}$ is a sequence of random permutations such that for all $n, \sigma_{n}$ is conjugacy invariant of size $n$.
Furthermore, we suppose that $\frac{\#_{1} \sigma_{n}}{n} \rightarrow \alpha$

## Theorem (Kim and Lee 2020)

$\frac{\operatorname{card}\left(D\left(\sigma_{n}\right)\right)-\frac{\left(1-\alpha^{2}\right) n}{2}}{\sqrt{n}} \xrightarrow[n \rightarrow \infty]{d} \mathscr{N}\left(0, \frac{1-4 \alpha^{3}+3 \alpha^{4}}{12}\right)$.
Goal: prove similar results for other functions.

## Classical Pattern

Let $\pi$ be a permutation of size $k$. An occurrence of the (classical) pattern $\pi$ in a permutation $\sigma$ is a vector $\left(i_{1}, \cdots, i_{k}\right)$ with $i_{1}<\cdots<i_{k}$ such that $\sigma\left(i_{1}\right) \ldots \sigma\left(i_{k}\right)$ has the same relative order as the elements of $\pi$. Examples:

- For the permutation $\sigma=2173456$, the vector $\left(i_{1}, i_{2}, i_{3}\right)=(2,3,7)$ is an occurrence of the pattern $\pi=132$ (176 has the same relative order as $\pi=132$.)
- An occurrence of 21 is an inversion.
- An occurrence of $123 \cdots k$ is an increasing subsequence of length $k$.


## Vincular Pattern

## Definition

Let $\pi$ be a permutation of size $k$ and $A$ be a subset of $[k-1]$. An occurrence of the vincular pattern $(\pi, A)$ in a permutation $\sigma$ is a vector $\left(i_{1}, \cdots, i_{k}\right)$ with $i_{1}<\cdots<i_{k}$ satisfying:

- $\left(i_{1}, \cdots, i_{k}\right)$ is an occurrence of the classical pattern $\pi$ in $\sigma$.
- For every $s$ in $A, i_{s+1}=i_{s}+1$.

Examples:

- $(\pi, \varnothing)$ : is the classical pattern $\pi$
- An occurrence of $(21,\{1\})$ : is a descent
- For the permutation $\sigma=2173456$, the vector $\left(i_{1}, i_{2}, i_{3}\right)=(2,3,7)$
- is an occurrence of the pattern $(\pi=132, A=\{1\})$
- not an occurrence of ( $\pi=132, A=\{1,2\}$ )

Notation: $\mathfrak{N}^{\pi, A}(\sigma)$ : pattern counts (number of occurrences of the patterns).

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## Uniform case

Fix $\Pi=(\pi, A)$, and let $k$ be the size of $\pi$.

## Theorem (Hofer (2018))

We assume that $\sigma_{n}$ uniform of size $n$

$$
\frac{\mathfrak{N}^{\Pi}\left(\sigma_{n}\right)-\mathbb{E}\left(\mathfrak{N}^{\Pi}\left(\sigma_{n}\right)\right)}{n^{k-\frac{1}{2}-\operatorname{card}(A)}} \underset{n \rightarrow \infty}{d} \mathscr{N}\left(0, \sigma_{\Pi}^{2}\right) .
$$

With

- $\sigma_{\Pi}^{2}>0$.

Generalises:

- $k=2$ : Fulman (2004)
- Consecutive: Goldstein (2005)
- Monotone: Bonà (2010)
- Classical: Janson et al. (2015)
- Without positivity: Féray (2013)


## Ewens

Recall: Ewens distribution.

$$
\mathbb{P}\left(\sigma_{n}=\sigma\right)=\frac{\theta^{\# \sigma}}{C_{n, \theta}}
$$

Fix $\Pi=(\pi, A)$, and $\theta \geq 0$. Let $k$ be the size of $\pi$.

## Theorem (Féray (2013))

We assume that $\sigma_{n}$ follows the Ewens distribution with parameter $\theta$. Then,

$$
\frac{\mathfrak{N}^{\Pi}\left(\sigma_{n}\right)-\mathbb{E}\left(\mathfrak{N}^{\Pi}\left(\sigma_{n}\right)\right)}{n^{k-\frac{1}{2}-\operatorname{card}(A)}} \underset{n \rightarrow \infty}{d} \mathscr{N}\left(0, \sigma_{\Pi}^{2}\right) .
$$

## Few cycles

Let $\sigma_{n}$ is conjugacy invariant of size $n$

## Theorem (Kammoun 2020)

We assume that $\frac{\#\left(\sigma_{n}\right)}{\sqrt{n}} \xrightarrow[n \rightarrow \infty]{d} 0$.
Then, $\xrightarrow[n^{k-\frac{1}{2}-\operatorname{card}(A)}]{\mathfrak{N}^{\Pi}\left(\sigma_{n}\right)} \underset{n \rightarrow \infty}{d} \mathscr{N}\left(0, \sigma_{\Pi}^{2}\right)$.

## Theorem (Hamaker and Rhoades (2022))

We assume that: for all $i \#_{i}\left(\sigma_{n}\right) \xrightarrow[n \rightarrow \infty]{d} 0$.
Then, $\frac{\left.\mathfrak{n}^{\Pi}\left(\sigma_{n}\right)-\mathbb{E} \mathfrak{N}^{\mathrm{\Pi}}\left(\sigma_{n}\right)\right)}{n^{k-\frac{1}{2}-\operatorname{card}(A)}} \xrightarrow[n \rightarrow \infty]{d} \mathscr{N}\left(0, \sigma_{\Pi}^{2}\right)$
If we combine both techniques.

## Theorem (Not written anywhere)

We assume that: for all $i \frac{\#_{i}\left(\sigma_{n}\right)}{\sqrt{n}} \xrightarrow[n \rightarrow \infty]{d} 0$.
Then, $\xrightarrow[n^{k-\frac{1}{2}-\operatorname{card}(A)}]{\left.\mathfrak{N}^{\Pi}\left(\sigma_{n}\right) \mathbb{E} \mathscr{N}^{\mathrm{\Pi}}\left(\sigma_{n}\right)\right)} \xrightarrow[n \rightarrow \infty]{d} \mathscr{N}\left(0, \sigma_{\Pi}^{2}\right)$

## Our result

Fix $\Pi=(\pi, A)$,

## Theorem (Féray and Kammoun (2023))

We assume that $\sigma_{n}$ is conjugacy invariant of size $n$ and that $\frac{\#_{1}\left(\sigma_{n}\right)}{n} \xrightarrow[n \rightarrow \infty]{d} \alpha$, $\frac{\#_{2}\left(\sigma_{n}\right)}{n} \xrightarrow[n \rightarrow \infty]{d} \beta$. Then

$$
\frac{\mathfrak{i}^{\Pi}\left(\sigma_{n}\right)-\mathbb{E}\left(\mathfrak{N}^{\Pi}\left(\sigma_{n}\right)\right)}{n^{k-\frac{1}{2}-\operatorname{card}(A)}} \underset{n \rightarrow \infty}{d} \mathscr{N}\left(0, \sigma_{\Pi, \alpha \beta}^{2}\right) .
$$

Moreover, if $A=\varnothing$, then $\sigma_{\Pi, \alpha, \beta}^{2}=0$ if and only if $(\alpha, \beta)=(1,0)$.
Remarks:

- Hofer (2018) implies that $\sigma_{\Pi, 0,0}^{2}>0$ for any $\Pi$.
- It is easy to see that $\sigma_{\Pi, 1,0}^{2}=0$ for any $\Pi$. (Identity)
- $\sigma_{\Pi, \alpha \beta}^{2}$ is a polynomial in $(\alpha \& \beta$ ). (Hamaker and Rhoades (2022))
- Dubach (2024) proved the same result for classical patterns $(A=\varnothing)+$ speed of convergence.
Conjecture: for any $\Pi, \sigma_{\Pi, \alpha, \beta}^{2}=0$ if and only if $(\alpha, \beta)=(1,0)$.
Questions: for which patterns, $\sigma_{\Pi, \alpha, \beta}^{2}$ does not depend on $\beta$ ? (consecutive)?


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## Comparison techniques

- Initially for the longest increasing subsequence / RSK (Kammoun 2018).
- Works for other combinatorial structures (coloured permutations, k-arrangements, etc.)
We give the proof of


## Theorem (Kammoun 2020)

We assume that $\frac{\#\left(\sigma_{n}\right)}{\sqrt{n}} \xrightarrow[n \rightarrow \infty]{d} 0$.
Then, $\xrightarrow[n^{k-\frac{1}{2}-\operatorname{card}(A)}]{\mathfrak{N}^{\Pi}\left(\sigma_{n}\right)} \underset{n \rightarrow \infty}{d} \mathscr{N}\left(0, \sigma_{\Pi}^{2}\right)$.

Simple random walk a directed version of the Cayley graph of $\mathfrak{S}_{n}$.


- If we start from any conjugacy invariant measure, the stationary measure is Ewens with parameter 0 .
- In each step, $\mathfrak{N}^{\mathrm{\Pi}}$ varies at most by $\frac{2}{k!} n^{k-\operatorname{card}(A)-1}$.

$$
\begin{aligned}
\left|\mathfrak{N}^{\Pi}\left(\sigma_{n}\right)-\mathfrak{N}^{\Pi}\left(\sigma_{n}^{u n i f}\right)\right| & \leq\left|\mathfrak{N}^{\Pi}\left(\sigma_{n}\right)-\mathfrak{N}^{\Pi}\left(\sigma_{0, n}^{E w}\right)\right|+\left|\mathfrak{N}^{\Pi}\left(\sigma_{0, n}^{E w}\right)-\mathfrak{N}^{\Pi}\left(\sigma_{n}^{u n i f}\right)\right| \\
& \leq \frac{2}{k!} n^{k-\operatorname{card}(A)-1}(\# \sigma_{n}+\underbrace{\# \sigma_{n}^{u n i f}}_{\approx \log (n)})
\end{aligned}
$$

We want that $\left|\mathfrak{N}^{\mathrm{I}}\left(\sigma_{n}\right)-\mathfrak{N}^{\mathrm{\Pi}}\left(\sigma_{n}^{u n i f}\right)\right|=o\left(n^{k-\operatorname{card}(A)-\frac{1}{2}}\right)$.
It is sufficient that $\# \sigma_{n}=o(\sqrt{n})$.

## Weighted dependency graphs

Initially developed by Féray (2018).
Works for other combinatorial structures.
We give a proof of

## Theorem (Féray and Kammoun (2023))

We assume that $\sigma_{n}$ is conjugacy invariant of size $n$ and that $\frac{\#_{1}\left(\sigma_{n}\right)}{n} \frac{d}{n \rightarrow \infty} \alpha$, $\frac{\#_{2}\left(\sigma_{n}\right)}{n} \xrightarrow[n \rightarrow \infty]{d} \beta$. Then

$$
\frac{\left.\mathfrak{N}^{\Pi}\left(\sigma_{n}\right)-\mathbb{E} \mathfrak{N}^{\Pi}\left(\sigma_{n}\right)\right)}{n^{k-\frac{1}{2}-\operatorname{card}(A)}} \xrightarrow[n \rightarrow \infty]{d} \mathscr{N}\left(0, \sigma_{\Pi, \alpha \beta}^{2}\right) .
$$

## Cumulants

## Definition

$$
\kappa_{r}\left(X_{1}, \ldots, X_{r}\right)=\left[t_{1} t_{2} \cdots t_{r}\right] \log \left(\mathrm{E}\left(\mathrm{e}^{\sum_{j=1}^{n} t_{j} X_{j}}\right)\right)
$$

For simplicity, we write $\kappa_{r}(X):=\kappa_{r}(X, \cdots, X)$.

- $X \sim \mathscr{N}\left(m, \sigma^{2}\right)$ if and only if for all $r \geq 3, \kappa_{r}(X)=0$
- If $X_{1}$ and $X_{2}$ are independent, then $\kappa_{r}\left(X_{1}+X_{2}\right)=\kappa_{r}\left(X_{1}\right)+\kappa_{r}\left(X_{2}\right)$
- $\kappa_{r}(X+C)=\kappa_{r}(X)$ if $r \geq 2$
- $\kappa_{r}(\alpha X)=\alpha^{r} \kappa_{r}(X)$
- If $\left\{X_{1}, \ldots X_{i}\right\}$ and $\left\{Y_{i+1}, \ldots Y_{r}\right\}$ are independent (and non-empty), then $\kappa_{r}\left(X_{1}, \ldots, X_{i}, Y_{i+1}, \ldots, Y_{r}\right)=0$
Proof of the CLT For $r \geq 3$

$$
K_{r}\left(\frac{\sum_{i=1}^{n} X_{i}-n \mathbb{E}\left(X_{1}\right)}{\sqrt{n}}\right)=K_{r}\left(\frac{\sum_{i=1}^{n} X_{i}}{\sqrt{n}}\right)=\frac{1}{n^{\frac{r}{2}}} \sum_{i=1}^{n} \kappa_{r}\left(X_{i}\right)=\frac{n}{n^{\frac{r}{2}}} \kappa_{r}\left(X_{1}\right)=o(1)
$$

## Weak dependency

- If $\left\{X_{1}, \ldots, X_{r}\right\}$ are "weakly dependent", then $\kappa_{r}\left(X_{1}, \ldots, X_{r}\right) \approx 0$.
- Dependency graphs: a graph with weights on the edges. Vertices are indexed by random variables, and weights measure the "dependency".
- If the weights are sufficiently "small", we have a CLT for the sum of the variables.


## Uniform Permutation

- Example: $\sigma_{n}$ is uniform and $A_{i, j}=1\left[\sigma_{n}(i)=j\right]$.
- If $i \neq j$ and $k \neq m$, then

$$
\mathbb{E}\left(A_{i, k} A_{j, m}\right)=\frac{1}{n(n-1)} \approx \frac{1}{n^{2}}=\mathbb{E}\left(A_{i, k}\right) \mathbb{E}\left(A_{j, m}\right)
$$

- if $k \neq m$, then $\mathbb{E}\left(A_{i, k} A_{i, m}\right)=0$ and $\mathbb{E}\left(A_{i, k}\right) \mathbb{E}\left(A_{j, m}\right)=\frac{1}{n^{2}}$.

For any $U=\left(i_{\ell}, j_{\ell}\right)_{1 \leq \ell \leq r}$, let $G(U)$, be the complete graph with vertices $U$ and the weight of $((i, j),(k, l))$ is $\begin{cases}1 & \text { if } i=k \text { or } j=l \\ \frac{1}{n} & \text { otherwise. }\end{cases}$

For example, if $U=((1,4),(1,2),(4,3),(1,2)), G(U)$


## Uniform Permutation

## Theorem (Féray 2018)

For all $r \geq 1$, there exists $C_{r}$ such that: For all integers $n$, for all $U=\left(i_{\ell}, j_{\ell}\right)_{1 \leq \ell \leq r}$

$$
\kappa_{r}\left(A_{i_{1}, j_{1}}, \ldots, A_{i_{r}, j_{r}}\right) \leq C_{r} \mathrm{M}(\mathrm{U}) n^{-\operatorname{card}(U)}
$$

where

- $M(U)$ is the maximum weight of a spanning tree of $G(U)$.
- $\operatorname{card}(U)$ is the number of distinct elements in $U$.

For example, if $U=((1,4),(1,2),(4,3),(1,2)), G(U)=$


For all $\left.n, \kappa_{r}\left(A_{1,4}, A_{1,2}, A_{4,3}, A_{1,2}\right)\right) \leq C_{4} \frac{1}{n} n^{-3}=C_{4} n^{-4}$

## New graphs

- $G^{1}(U)$, the complete graph with vertices $U$ and the weight of $((i, j),(k, l))$ is 1 if $i=k$ or $j=l$ or $i=j$ or $k=l$, and $\frac{1}{n}$ otherwise.

For example, if $U=((1,4),(1,2),(4,3),(1,2)), G^{1}(U)=$


- $G^{2}(U):=([n], E=U)$

For example, if $U=((1,4),(1,2),(4,3),(1,2)), G^{2}(U)=1$


## Uniform Permutation within a Conjugacy Class

$\sigma_{n}^{\lambda}$ is uniform within the conjugacy class $\lambda$ and $A_{i, j}=1\left[\sigma_{n}^{\lambda}(i)=j\right]$.

## Theorem (Féray and Kammoun 2023)

For all $r \geq 1$, there exists $C_{r}$ such that: For all integers $n$, for all $U=\left(i_{\ell}, j_{\ell}\right)_{1 \leq \ell \leq r}$

$$
\kappa_{r}\left(A_{i_{1}, j_{1}}, \ldots, A_{i_{r}, j_{r}}\right) \leq C_{r} \mathrm{M}(\mathrm{U}) n^{C C(U)-\operatorname{card}(U)}
$$

where

- $M(U)$ is the maximum weight of a spanning tree of $G^{1}(U)$, the complete graph with vertices $U$ and the weight of $((i, j),(k, l))$ is 1 if $i=k$ or $j=l$ or $i=j$ or $k=l$, and $\frac{1}{n}$ otherwise.
- $\operatorname{card}(U)$ is the number of distinct elements in $U$.
- $C C(U)$ the number of nontrivial connected components in the graph $G^{2}(U)=([n], E=U)$


## Application: Patterns

If we denote by $X^{(\pi, A)}$ the number of occurrences of the pattern $(\pi, A)$, we have

$$
X^{(\pi, A)}\left(\sigma_{n}^{\lambda}\right)=\sum_{\substack{i_{1} \leq \cdots \cdots i_{k} \\ i_{s+1}=i_{s}+1 \text { or sec } s \in A}} \sum_{\substack{j_{n}, \ldots j_{k}-j_{k}(1)<\cdots \cdots i_{n}-1(k)}} A_{i_{1}, j_{1}} \cdots A_{i_{k} j_{k}} .
$$

To conclude: The magic of weighted dependency graphs: We can "easily" move from controlling mixed cumulants of $\left\{A_{i, j}:(i, j) \in[n]^{2}\right\}$ to controlling mixed cumulants of $\left\{A_{i_{1}, i_{2}} \cdots A_{i_{r}, j_{r}}:\left(i_{1}, j_{1}, \ldots, i_{r}, j_{r}\right) \in[n]^{2 r}\right\}$. We obtain

$$
\kappa_{r}\left(X^{(\pi, A)}\left(\sigma_{n}^{\lambda}\right)\right) \leq C_{k, r} n^{r(k-\operatorname{card}(A)-1)+1}
$$

and thus

$$
\kappa_{r}\left(\frac{X^{(\pi, A)}\left(\sigma_{n}^{\lambda}\right)-\mathbb{E}\left(X^{(\pi, A)}\left(\sigma_{n}^{\lambda}\right)\right)}{n^{k-\operatorname{card}(A)-\frac{1}{2}}}\right) \leq C_{k, r} n^{1-\frac{r}{2}}
$$

## Motivation: universality

## Central Limit Theorem

Let $X_{1}, X_{2}, \ldots, X_{n}$ be i.i.d with $\operatorname{Var}\left(X_{i}\right)=\sigma^{2}<+\infty$. Then,

$$
\sqrt{n}\left(\frac{1}{n} \sum_{i=1}^{n} X_{i}-\mathbb{E}\left(X_{1}\right)\right) \xrightarrow{d} \mathscr{N}\left(0, \sigma^{2}\right)
$$

The limit is universal (does not depend on the distribution of $X_{i}$ ).

Symmetry/independence + control $=$ universality

## Fisher-Tippett-Gnedenko Theorem

Let $X_{1}, X_{2}, \ldots, X_{n}$ be i.i.d and $M_{n}=\max \left(X_{1}, X_{2}, \ldots, X_{n}\right)$.
Suppose there exist constants $a_{n}>0$ and $b_{n}$ such that, for every real $x$,

$$
\mathbb{P}\left(\frac{M_{n}-b_{n}}{a_{n}} \leq x\right) \rightarrow G(x)
$$

where $G(x)$ is a non-degenerate cumulative distribution function. Then, $G$ is the cumulative distribution function of a Gumbel, Fréchet, or Weibull variable.
The limit fluctuations depend on the tail of the distribution of $X_{1}$.

Symmetry/Independence + Control $=$ Universality

## Wigner Matrices

Let's define the symmetric matrix $M$ as

$$
M=\frac{1}{\sqrt{n}}\left[\begin{array}{ccccc}
a_{1,1} & a_{1,2} & \ldots & \ldots & a_{1, n} \\
a_{1,2} & a_{2,2} & \ldots & \ldots & a_{2, n} \\
\vdots & \vdots & \ddots & \vdots & \vdots \\
\vdots & \vdots & \vdots & \ddots & \vdots \\
a_{1, n} & a_{2, n} & \ldots & \ldots & a_{n, n}
\end{array}\right]
$$

The entries $\left\{a_{i, j}\right\}_{1 \leq i \leq j \leq n}$ are i.i.d. such that $\mathbb{E}\left(a_{1,1}\right)=0$ and $\mathbb{E}\left(a_{1,1}^{2}\right)=1$.
Let $\lambda_{1} \leq \lambda_{2} \leq \ldots \leq \lambda_{n}$ be the eigenvalues of $M$.

## Histogram of Eigenvalues



Gaussian entries

1 avec proba 0.5 et -1 avec proba 0.5


Entries 1 or - 1

## Wigner's theorem

"The histogram of eigenvalues is not far from a semi-circle"

## Theorem

The empirical spectral measure of the eigenvalues of $M$

$$
\frac{1}{n} \sum_{i=1}^{n} \delta_{\lambda_{i}},
$$

converges weakly to the semi-circular law of Wigner as $n$ tends to infinity.

But also ${ }^{*}$,

- The largest eigenvalue converges to 2
- The fluctuations of the largest eigenvalue are of Tracy-Widom type
- Large deviations of the largest eigenvalues are universal
- The joint limit fluctuations of the first $k$ eigenvalues are universal
- The local limit laws are universal
- The fluctuations of the number of points in [a,b] are universal

And for random permutations?
*Some conditions apply on the moments / the tail of the distribution

## Longest Decreasing Subsequence

- $\left(\sigma\left(i_{1}\right), \ldots, \sigma\left(i_{k}\right)\right)$ is a decreasing subsequence of $\sigma$ if $i_{1}<i_{2}<\cdots<i_{k}$ and $\sigma\left(i_{1}\right)>\cdots>\sigma\left(i_{k}\right)$.


## Longest Decreasing Subsequence

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- $\operatorname{LDS}(\sigma)$ : The length of the longest decreasing subsequence of $\sigma$.


## Longest Decreasing Subsequence

- $\left(\sigma\left(i_{1}\right), \ldots, \sigma\left(i_{k}\right)\right)$ is a decreasing subsequence of $\sigma$ if $i_{1}<i_{2}<\cdots<i_{k}$ and $\sigma\left(i_{1}\right)>\cdots>\sigma\left(i_{k}\right)$.
- $\operatorname{LDS}(\sigma)$ : The length of the longest decreasing subsequence of $\sigma$.
- Example:

$$
\sigma=\left(\begin{array}{llllllll}
1 & 2 & 3 & 4 & 5 & 6 & 7 & 8 \\
6 & 1 & 8 & 7 & 5 & 2 & 4 & 3
\end{array}\right)
$$

$\operatorname{LDS}(\sigma)=5$.

## Longest Decreasing Subsequence: Universality

We assume that $\sigma_{n}$ is conjugation invariant and $\frac{\#_{1}\left(\sigma_{n}\right)}{n} \rightarrow \alpha$

## Theorem (Dubach (2024+))

$$
\xrightarrow[{\sqrt{n}}]{\operatorname{LDS}\left(\sigma_{n}\right)} \xrightarrow[n \rightarrow \infty]{d} 2 \sqrt{1-\alpha}
$$

## Theorem (Kammoun 2018)

If
$n^{\frac{-1}{6}} \min _{1 \leq i \leq n}\left(\left(\sum_{j=1}^{i} \#_{j}\left(\sigma_{n}\right)\right)+\frac{\sqrt{n}}{i} \sum_{j=i+1}^{n} \#_{j}\left(\sigma_{n}\right)\right) \xrightarrow[n \rightarrow \infty]{\mathbb{巴}}$
0, then, $\frac{\operatorname{LDS}\left(\sigma_{n}\right)-2 \sqrt{n}}{\sqrt[b]{n}} \xrightarrow[n \rightarrow \infty]{d}$ Tracy Widom

## Theorem (Guionnet, Kammoun 2023)

If $\sigma_{n}$ is conjugacy invariant and $\#\left(\sigma_{n}\right)=o(\sqrt{n})$. Then, $\frac{\operatorname{LDS}\left(\sigma_{n}\right)}{\sqrt{n}}$ satisfies a LD principle

- with speed $\sqrt{n}$ and rate function $J_{L D S, \frac{1}{2}}$.
- with speed $n$ and rate function $J_{L D S, 1}$

With,

$$
\begin{gathered}
J_{L D S, \frac{1}{2}}(x)=\left\{\begin{array}{ll}
2 x \cosh ^{-1} \frac{x}{2} & \text { if } x>2 \\
+\infty & \text { if } x \leq 2 .
\end{array} .\right. \\
J_{L D S, 1}(x)=\left\{\begin{array}{ll}
-1+\frac{x^{2}}{4}+2 \ln \left(\frac{x}{2}\right)-\left(2+\frac{x^{2}}{2}\right) \ln \left(\frac{2 x^{2}}{4+x^{2}}\right) & \text { if } 0<x \leq 2 \\
0 & \text { if } x>2 \\
+\infty & \text { if } x \leq 0
\end{array} .\right.
\end{gathered}
$$

In other words: if $\sigma_{n}$ is conjugation invariant and \#( $\sigma$ ) "is low" then
$-\log \left(\mathbb{P}\left(\frac{\operatorname{LDS}\left(\sigma_{n}\right)}{\sqrt{n}} \approx x\right)\right) \approx\left\{\begin{array}{ll}\left(-1+\frac{x^{2}}{4}+2 \ln \left(\frac{x}{2}\right)-\left(2+\frac{x^{2}}{2}\right) \ln \left(\frac{2 x^{2}}{4+x^{2}}\right)\right) n & \text { if } x \in] 0,2[] \\ 2 x \cosh ^{-1}\left(\frac{x}{2}\right) \sqrt{n} & \text { if } x>2 \\ +\infty & \text { if } x \leq 0 \\ 0 & \text { if } x=2\end{array}\right.$.
The same phenomenon appears for $\lambda_{1}$ (Wigner Matrices).

## What we know

Type 1: Local events

- $\mathbb{P}(S \subset D(\sigma))$
- $\mathbb{P}(\sigma(10)>10)$

Type 2: LLN / first order / global convergence

- $\frac{\mathfrak{x}^{\mathrm{n}}}{n^{k-\operatorname{card}(A)}}$
- $\frac{\operatorname{LDS}}{\sqrt{n}}$

The limit depends only on $\frac{\#_{1}}{n}$

Type 3: fluctuations (Poisson / Normal)

- $\operatorname{Tr}\left(\left(\sigma_{n} \rho_{n} \pi_{n} \sigma_{n}^{-1} \rho_{n}^{-1} \pi_{n}\right)^{2024}\right)$
- $\frac{\mathfrak{\Re}^{\mathrm{H}}-\mathbb{E}\left(\mathfrak{N}^{\mathrm{H}}\right)}{n^{k-\operatorname{carar}(\hat{1})-\frac{1}{2}}}$

The limit depends on $\frac{\#_{1}}{n}$ and $\frac{\#_{2}}{n^{\alpha}}$ for some $\alpha$

Type 4: others

- $\frac{\text { LDS }-2 \sqrt{n}}{n^{\frac{1}{6}}}$
- Large deviations.

Universality if \# is low. There is still much work to be done.

|  | Exact calculation | Representations | Method of moments (and its variants) |  | Comparison | Geometric |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  |  |  | Random matrices | Dependency graphs |  |  |
| Universality Permutations | Fulman, Kim, Lee | Hamaker and Rhoades | Kammoun and Maïda | Féray and Kammoun | Guionnet and Kammoun | Dubach |
| Functions | Descents Valleys | Descents <br> Inversions <br> Partterns <br> (classic, <br> (bi)-vincular, LAS | Trace of words | Descents <br> Inversions <br> Partterns <br> (classic, <br> vincular) <br> Long.Altern.Sub | Descents <br> Inversions <br> Partterns <br> (classic, <br> vincular) <br> Long.Altern.Sub <br> LDS <br> Long.Com.Sub. <br> RSK Bord <br> RSK shape <br> Gran dev | Inversion <br> Partterns <br> classic <br> RSK Shape <br> LDS(order 1) |
| Limits | Normal | Constant | Poisson <br> Mixtures | Normal | Normal Tracy Widom Airy, VKLS | Normal |
| Arxiv | $\begin{array}{\|l} \hline 2018,2018 \\ 2019 \\ \hline \end{array}$ | 2022 | 2019, 2022 | 2023 | $\begin{aligned} & \hline 2018,2020 \\ & 2023 \\ & \hline \end{aligned}$ | 2024+ |


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## Merci de votre attention

