

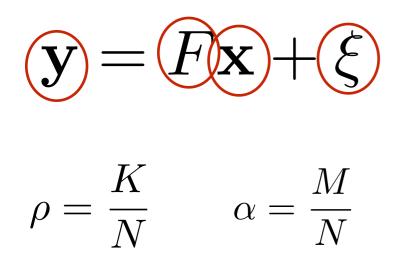
On convergence of Approximate Message Passing

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Compressed Sensing

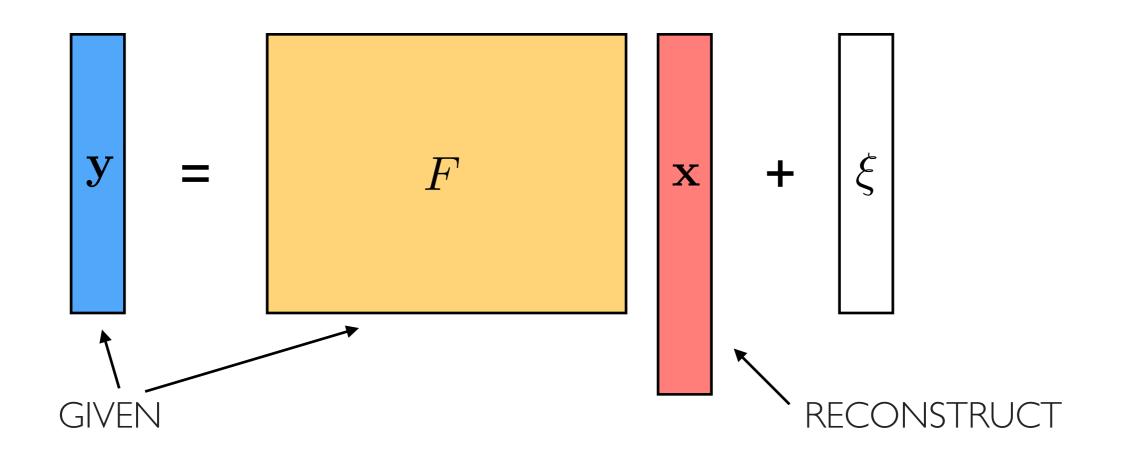


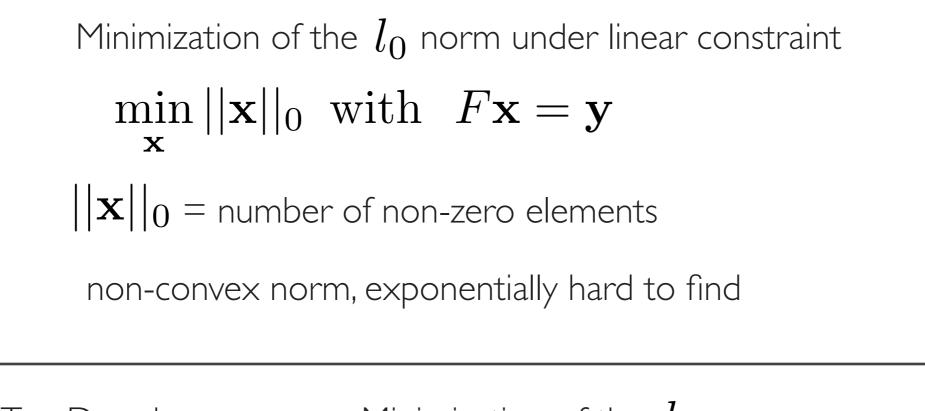
the signal is an N components vector only K<N components are non-zero

the measurement is an M<N components vector

MxN random matrix with i.i.d. elements

white noise with variance $\langle \xi^2
angle = \Delta$

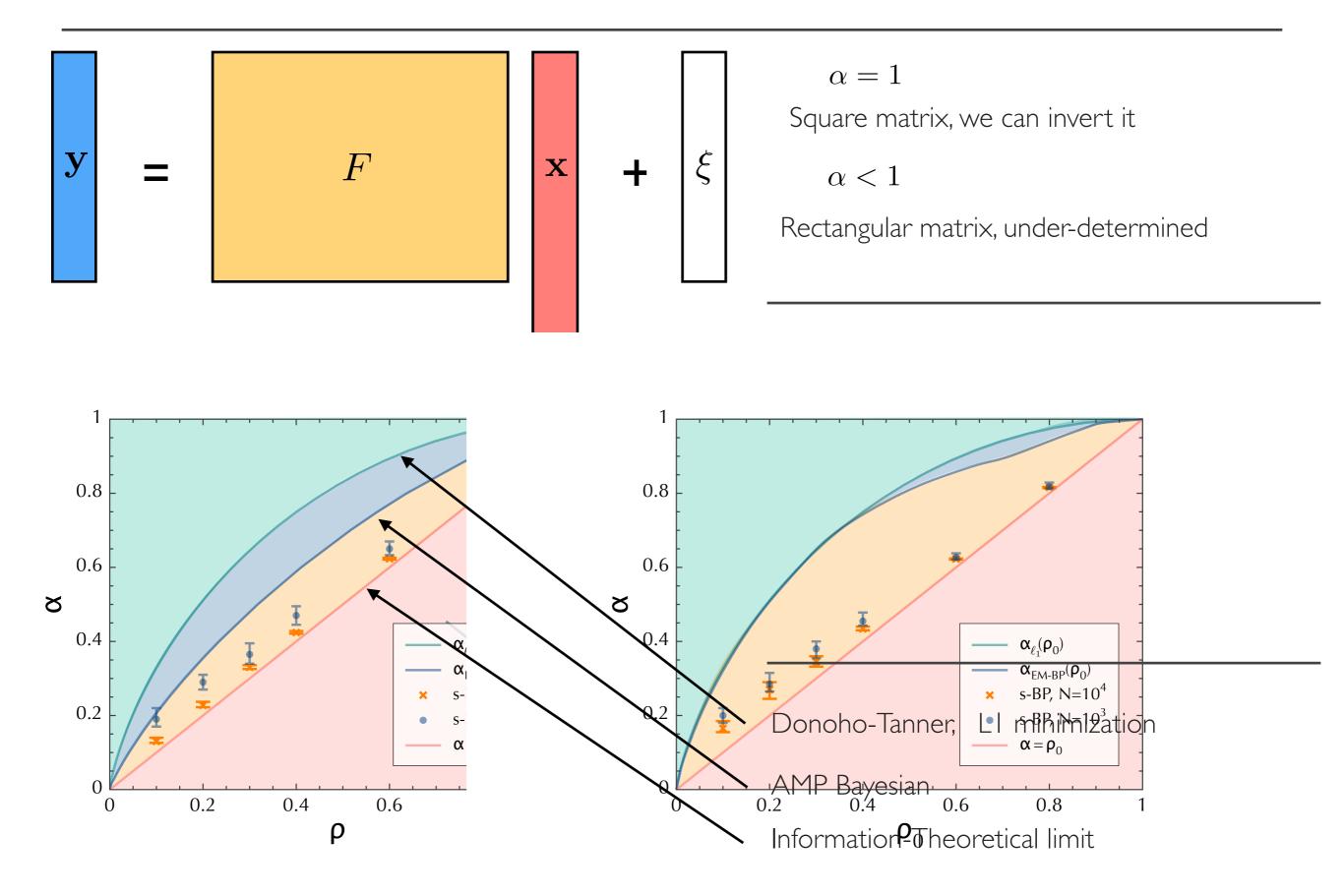




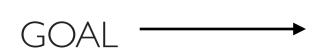
Candès, Tao, Donoho
$$\longrightarrow$$
 Minimization of the l_1 norm
 $||\mathbf{x}||_1 = \sum_{i=1}^N |x_i|$ convex norm, easy to minimize

The l_1 norm well approximates the l_0 norm

The Donoho-Tanner line



Bayesian setting



Reconstruct the signal, given the measurement vector, the measurement matrix and a prior knowledge of the (sparse) distribution of signal elements

Approximate Message Passing .

Donoho, Maleki, Montanari (2009)

Powerful algorithm. Convergence issues.

$$\mathbf{y} = F\mathbf{x} + \xi \qquad F_{\mu i} = \frac{\gamma}{N} + \frac{1}{\sqrt{N}}\mathcal{N}(0, 1) \qquad P(x) = (1 - \rho)\delta(x) + \rho\mathcal{N}(0, 1)$$

Simplest case in which Approximate Massage Passing (AMP) has convergence problems.

If the mean is sufficiently large then AMP displays violent divergencies.

This kind of divergencies are observed in many other cases and are the main obstacle to a wide use of AMP.

In this simple case there are workarounds that ensure convergence, like a "mean-removal" procedure. BUT it is interesting because want to understand the origin of the nonconvergence that, we argue, is of the same nature in more complicated settings.

Bayes formula
$$\longrightarrow P(\mathbf{x}|\mathbf{F}, \mathbf{y}) = \frac{P(\mathbf{x}|\mathbf{F})P(\mathbf{y}|\mathbf{F}, \mathbf{x})}{P(\mathbf{y}|\mathbf{F})}$$

Conditional probability of the measurement vector $\longrightarrow P(\mathbf{y}|\mathbf{F}, \mathbf{x}) = \prod_{\mu=1}^{M} \frac{1}{\sqrt{2\pi\Delta_{\mu}}} e^{-\frac{1}{2\Delta_{\mu}}(y_{\mu} - \sum_{i=1}^{N} F_{\mu i}x_{i})^{2}}$

$$P(\mathbf{x}|\mathbf{F},\mathbf{y}) = \frac{1}{Z(\mathbf{y},\mathbf{F})} \prod_{i=1}^{N} \frac{[(1-\rho)\delta(x_i) + \rho\phi(x_i)]}{[(1-\rho)\delta(x_i) + \rho\phi(x_i)]} \prod_{\mu=1}^{M} \frac{1}{\sqrt{2\pi\Delta_{\mu}}} e^{-\frac{1}{2\Delta_{\mu}}(y_{\mu} - \sum_{i=1}^{N} F_{\mu i}x_i)^2}$$

If we know exactly the prior distribution on the signal elements and on the noise we are in the so-called BAYES OPTIMAL setting

In the following we will consider that this is the case. When it is not the case, the prior can be efficiently learned adding a step to the algorithm that I will present. (I will not talk about this)

Belief Propagation (Cavity method)

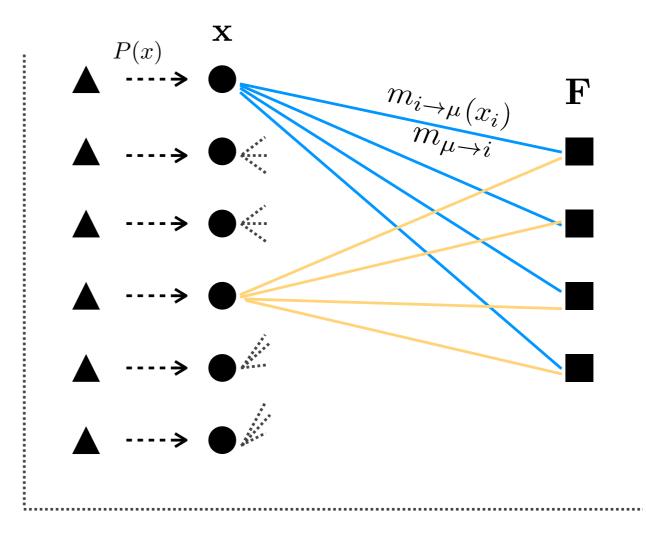
Two kinds of nodes: factors (matrix lines) and variables (signal elements)

We can introduce a third kind of nodes: the prior distribution on the signal elements, local field.

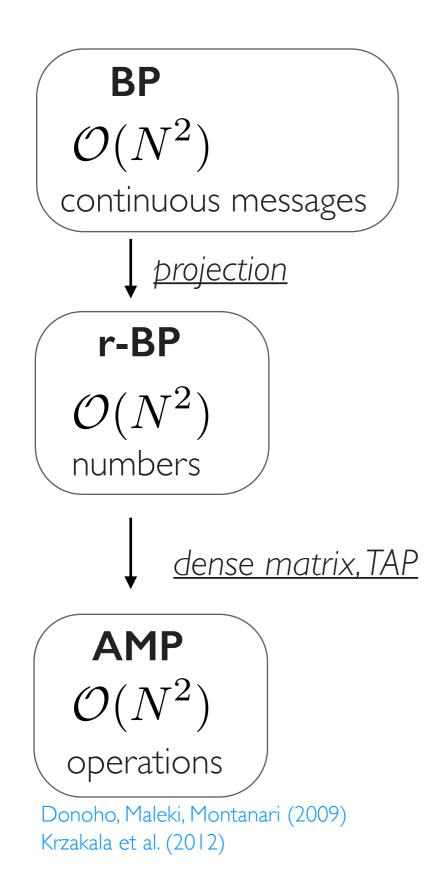
Belief propagation works for: locally tree-like graphs or densely and weakly connected graphs.

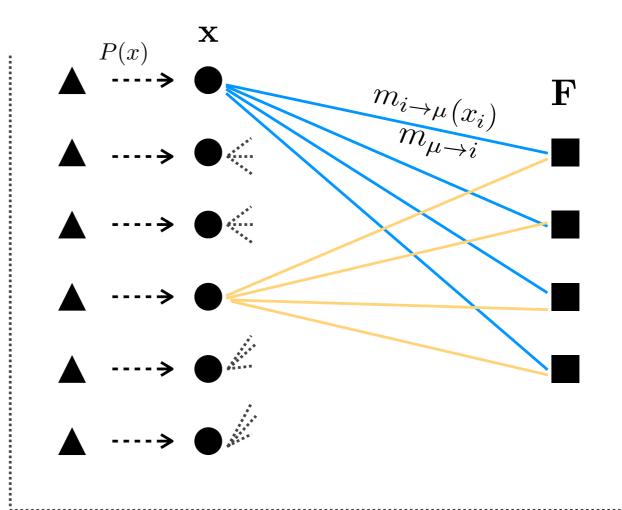
Messages represent an approximation to the marginal distribution of a variable.

Messages are updated according to a sequential or parallel schedule until convergence (fixed point).



Belief Propagation, r-BP and AMP





For the last step one assumes parallel update

In this case, fast matrix multiplication algorithms can be applied, reducing the complexity to

 $N\log(N)$

AMP Algorithm

$$\begin{split} V_{\mu}^{t+1} &= \sum_{i} F_{\mu i}^{2} v_{i}^{t}, \qquad (1) \\ \omega_{\mu}^{t+1} &= \sum_{i} F_{\mu i} a_{i}^{t} - \frac{(y_{\mu} - \omega_{\mu}^{t})}{\Delta + V_{\mu}^{t}} \sum_{i} F_{\mu i}^{2} v_{i}^{t}, \qquad (2) \\ (\Sigma_{i}^{t+1})^{2} &= \left[\sum_{\mu} \frac{F_{\mu i}^{2}}{\Delta + V_{\mu}^{t+1}} \right]^{-1}, \qquad (3) \\ R_{i}^{t+1} &= a_{i}^{t} + \frac{\sum_{\mu} F_{\mu i} \frac{(y_{\mu} - \omega_{\mu}^{t+1})}{\Delta + V_{\mu}^{t+1}}}{\sum_{\mu} \frac{F_{\mu i}^{2}}{\Delta + V_{\mu}^{t+1}}}, \qquad (4) \\ a_{i}^{t+1} &= f_{1} \left((\Sigma_{i}^{t+1})^{2}, R_{i}^{t+1} \right), \\ v_{i}^{t+1} &= f_{2} \left((\Sigma_{i}^{t+1})^{2}, R_{i}^{t+1} \right). \qquad (5) \\ v_{i}^{t+1} &= f_{2} \left((\Sigma_{i}^{t+1})^{2}, R_{i}^{t+1} \right). \qquad (6) \end{split}$$

 $f_k(\Sigma^2, R) \rightarrow$ k-th connected cumulants w.r.t. the measure

$$\mathcal{Q}(x) = \frac{1}{Z(\Sigma^2, R)} P(x) \frac{e^{-\frac{(x-R)^2}{2\Sigma^2}}}{\sqrt{2\pi\Sigma^2}}$$

 a_i and v_i are the AMP estimators for the mean and variance of the i-th signal component.

AMP Algorithm

$$\begin{split} V_{\mu}^{t+1} &= \sum_{i} F_{\mu i}^{2} v_{i}^{t}, \qquad (1) \\ \omega_{\mu}^{t+1} &= \sum_{i} F_{\mu i} a_{i}^{t} - \frac{(y_{\mu} - \omega_{\mu}^{t})}{\Delta + V_{\mu}^{t}} \sum_{i} F_{\mu i}^{2} v_{i}^{t}, \qquad (2) \\ (\Sigma_{i}^{t+1})^{2} &= \left[\sum_{\mu} \frac{F_{\mu i}^{2}}{\Delta + V_{\mu}^{t+1}} \right]^{-1}, \qquad (3) \\ R_{i}^{t+1} &= a_{i}^{t} + \frac{\sum_{\mu} F_{\mu i} \frac{(y_{\mu} - \omega_{\mu}^{t+1})}{\Delta + V_{\mu}^{t+1}}}{\sum_{\mu} \frac{F_{\mu i}^{2}}{\Delta + V_{\mu}^{t+1}}}, \qquad (4) \\ a_{i}^{t+1} &= f_{1} \left((\Sigma_{i}^{t+1})^{2}, R_{i}^{t+1} \right), \qquad (5) \\ v_{i}^{t+1} &= f_{2} \left((\Sigma_{i}^{t+1})^{2}, R_{i}^{t+1} \right). \qquad (5) \end{split}$$

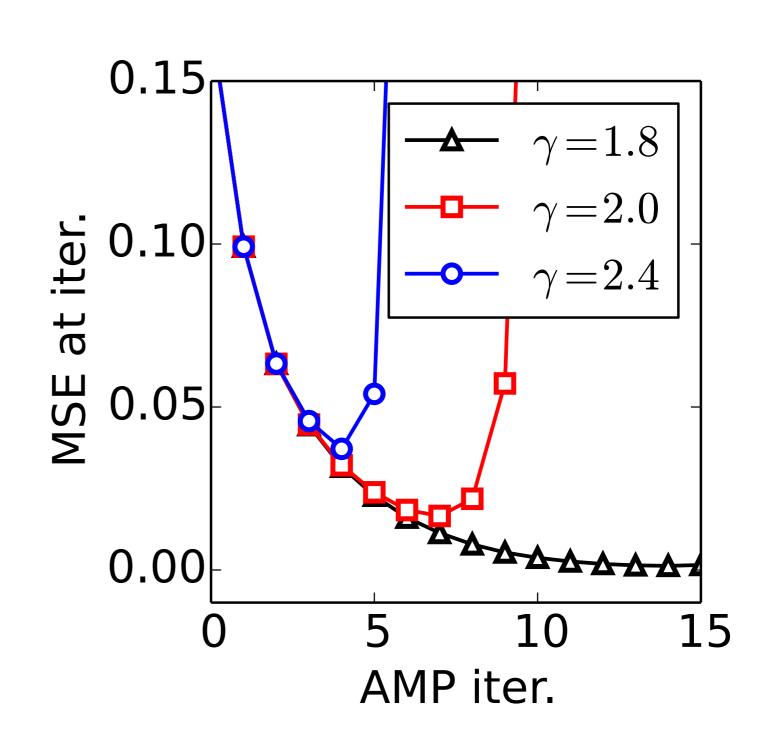
$$F_{\mu i} = \frac{\gamma}{N} + \frac{1}{\sqrt{N}}\mathcal{N}(0,1)$$

The AMP algorithm does NOT depend explicitly on the value of the mean of the matrix.

Bayes optimal case.

Given a certain (sufficiently high) measurement ratio.

Very small or zero noise.



State evolution is the asymptotic analysis of the average performance of the inference algorithm when the size of the signal goes to infinity.

It gives a good indication of what happens in a practical situation if the size of the signal is sufficiently large.

It can be obtained rigorously in simple cases and non rigorously with the replica method in more involved cases.

State Evolution (infinite N)

Bayati, Montanari (rigorous in the zero-mean case) '11 Krzakala et al. (replicas in the zero-mean case) '12 Caltagirone, Krzakala, Zdeborova (replicas in the non-zero-mean case) '14

$$E^{t} = \frac{1}{N} \sum_{i=1}^{N} (s_{i} - a_{i}^{t})^{2} \qquad V^{t} = \frac{1}{N} \sum_{i=1}^{N} v_{i} \qquad D^{t} = \frac{1}{N} \sum_{j} (s_{j} - a_{j}^{t})$$

$$V^{t+1} = \int \mathrm{d}s \, P(s) \, \int \mathcal{D}z \times f_2\left(\frac{\Delta + V^t}{\alpha}, s + z\mathcal{A}(E^t, D^t) + \gamma^2 D^t\right)$$

$$E^{t+1} = \int \mathrm{d}s \, P(s) \, \int \mathcal{D}z \times \left[s - f_1 \left(\frac{\Delta + V^t}{\alpha}, s + z \mathcal{A}(E^t, D^t) + \gamma^2 D^t \right) \right]^2$$

$$D^{t+1} = \int \mathrm{d}s \, P(s) \, \int \mathcal{D}z \times \left[s - f_1 \left(\frac{\Delta + V^t}{\alpha}, s + z \mathcal{A}(E^t, D^t) + \gamma^2 D^t \right) \right]$$

with
$$\mathcal{A}(E^t, D^t) = \sqrt{\frac{E^t + \Delta + \gamma^2 (D^t)^2}{\alpha}}$$

If the mean is zero the density evolution that does not depend on D

The Nishimori Condition

Bayes optimal setting
$$\longrightarrow \begin{array}{cc} E^t = V^t & D^t = 0 \end{array}$$

 $E^{t+1} = V^{t+1} & D^{t+1} = 0 \end{array}$

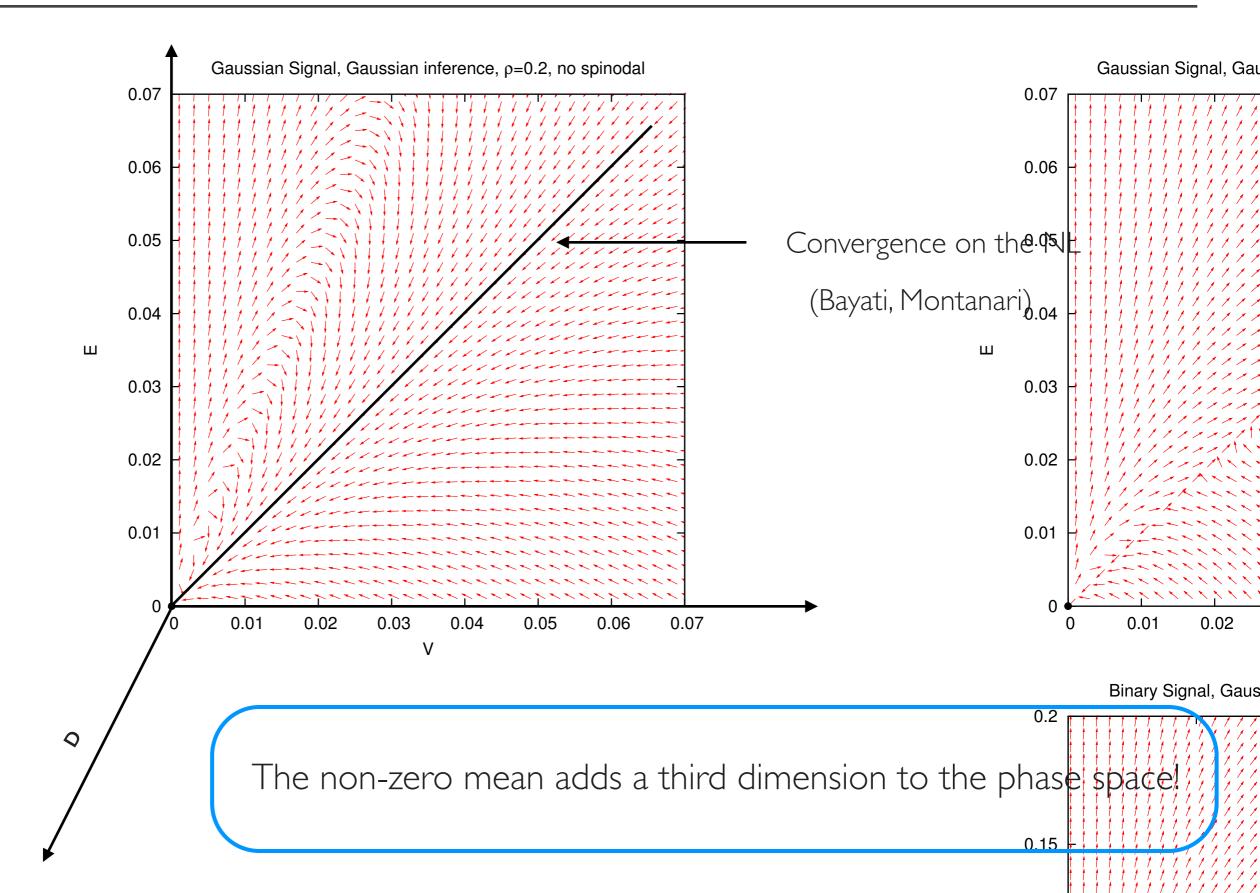
Therefore, analytically, if the evolution starts (exactly) on the Nishimori Line it stays on it until convergence.



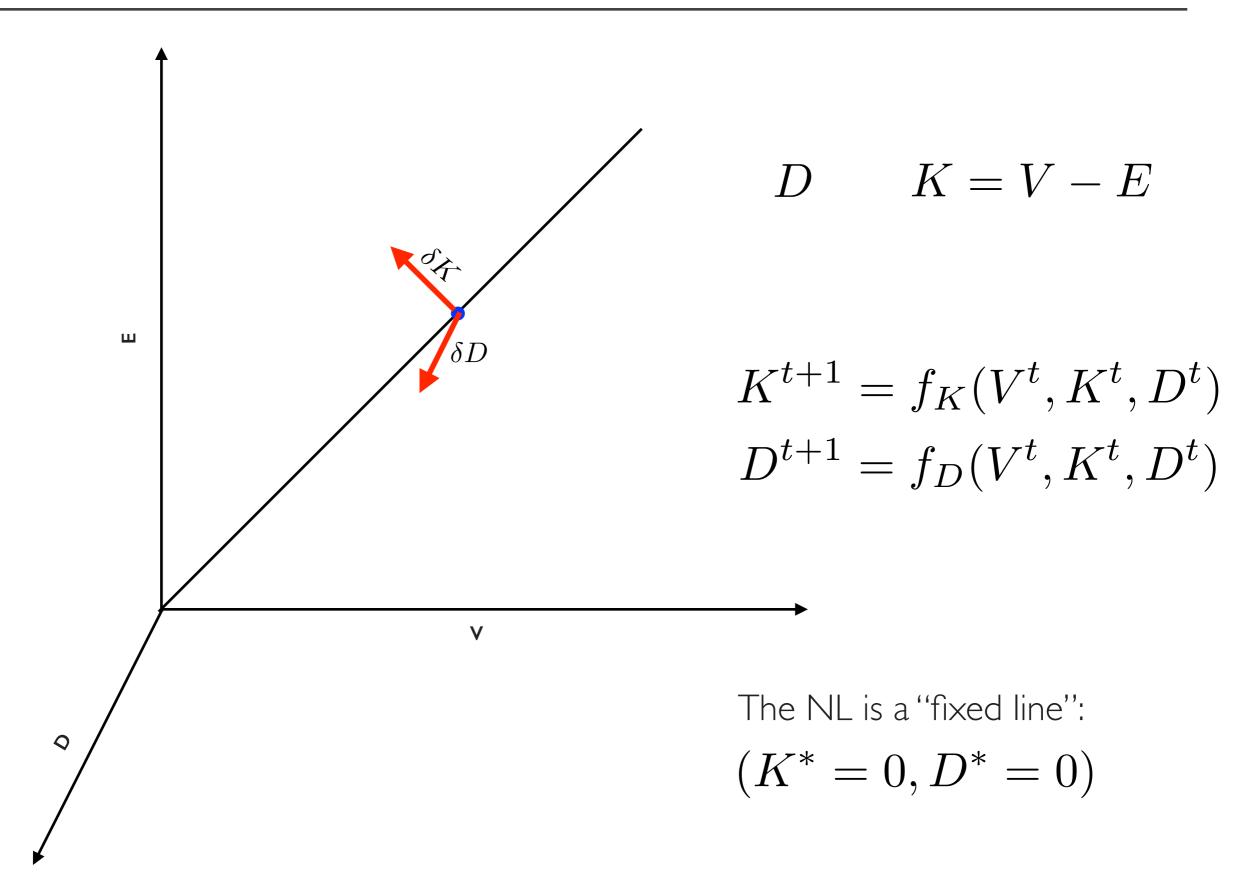
What is the effect of small perturbations with respect to the NL?

- Very small fluctuations due to numerical precision in the DE
- Fluctuations due to finite size in the AMP algorithm

Zero-mean case



Stability Analysis (I)



Stability Analysis (II)

We linearize the equations with

$$\frac{\delta K^t = K^t - K^*}{\delta D^t = D^t - D^*} \qquad \left(\begin{array}{c} \delta K^{t+1} \\ \delta D^{t+1} \end{array}\right) = \mathcal{M} \cdot \left(\begin{array}{c} \delta K^t \\ \delta D^t \end{array}\right)$$

$$\mathcal{M} = \begin{pmatrix} \partial_K f_K(V^t, 0, 0) & \partial_D f_K(V^t, 0, 0) \\ \partial_K f_D(V^t, 0, 0) & \partial_D f_D(V^t, 0, 0) \end{pmatrix}$$

Stability Analysis (II)

We linearize the equations with

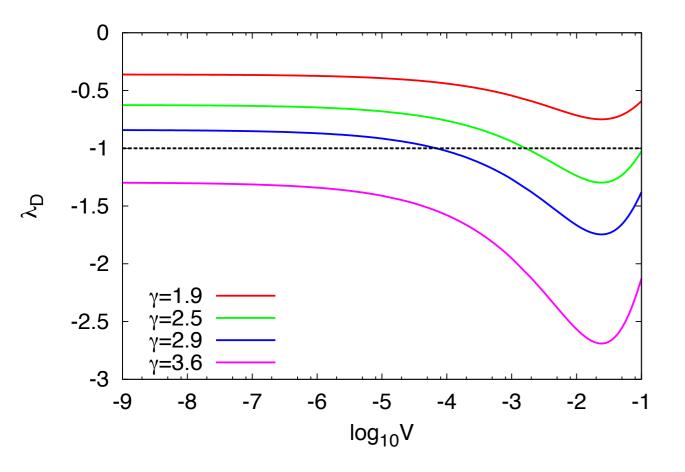
$$\frac{\delta K^t = K^t - K^*}{\delta D^t = D^t - D^*} \qquad \left(\begin{array}{c} \delta K^{t+1} \\ \delta D^{t+1} \end{array}\right) = \mathcal{M} \cdot \left(\begin{array}{c} \delta K^t \\ \delta D^t \end{array}\right)$$

$$\mathcal{M} = \begin{pmatrix} \partial_K f_K(V^t, 0, 0) & 0 \\ 0 & \partial_D f_D(V^t, 0, 0) \end{pmatrix}$$

When the signal is Gauss-Bernoulli with zero mean, the off-diagonal terms vanish.

$$\begin{aligned} \partial_D f_D(V^t) &= -\frac{\alpha \gamma^2}{\Delta + V^t} \int ds \, P(s) \int \mathcal{D}z \, f_2 \left(A^2, s + zA \right) = -\frac{\alpha \gamma^2 V^{t+1}}{\Delta + V^t} \,, \quad \overleftarrow{\Delta D} \\ \partial_K f_K(V^t) &= -\frac{1}{2} \frac{1}{\Delta + V^t} \int ds \, P(s) \int \mathcal{D}z \, \left\{ f_4 \left(A^2, s + zA \right) + 2(f_2 \left(A^2, s + zA \right))^2 \right. \\ &+ 2 \left[f_1 \left(A^2, s + zA \right) - s \right] f_3 \left(A^2, s + zA \right) \right\} \,, \end{aligned}$$

$$\rho = 0.1 \quad \alpha = 0.3 \quad \Delta = 10^{-10}$$

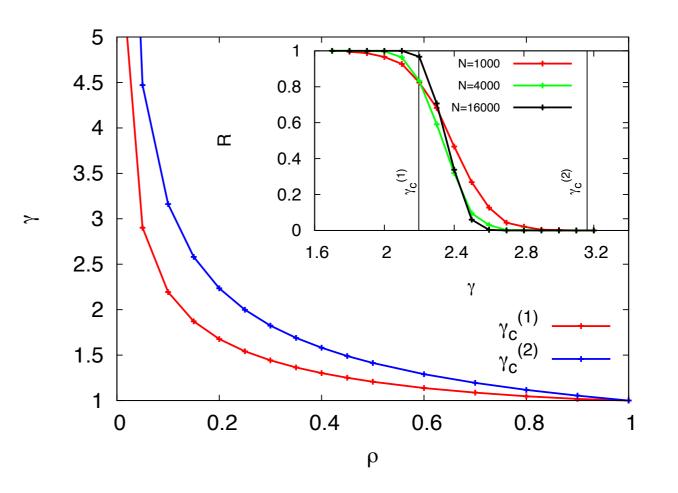


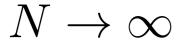
- $\gamma < \gamma_c^{(1)}$ the eigenvalue is always less than I in modulus.
- $\gamma_c^{(1)} < \gamma < \gamma_c^{(2)}$ the eigenvalue becomes larger than 1 in a limited region.
- $\gamma > \gamma_c^{(2)}$ the eigenvalue is larger than I in modulus down to the fixed point.

For zero measurement noise both the critical values do NOT depend on the undersampling rate $\,\alpha\,$.

For weak noise only the second critical value has a very weak dependence on both Δ and α .

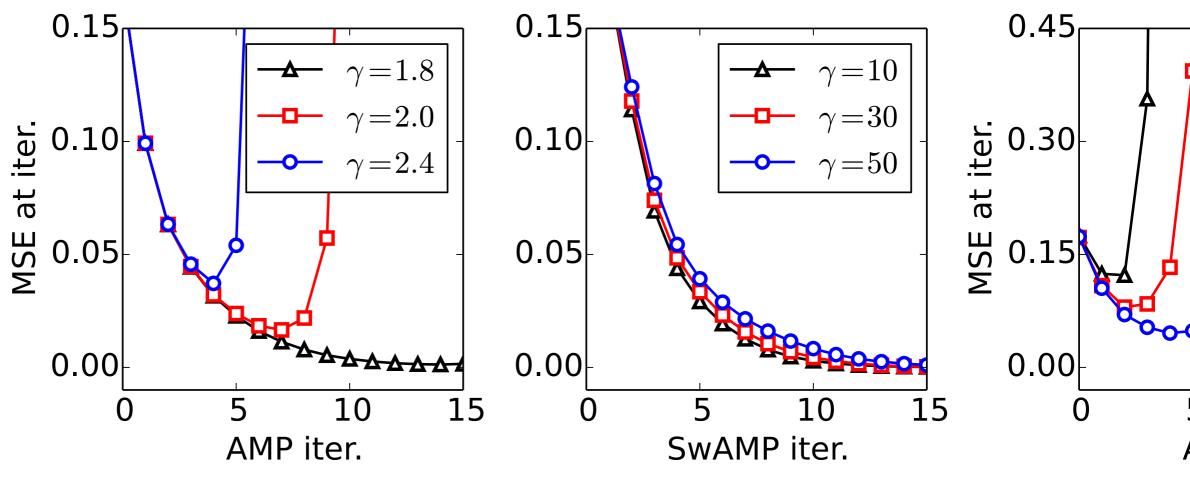
[Inset] Convergence Rate of the AMP algorithm for different signal sizes.



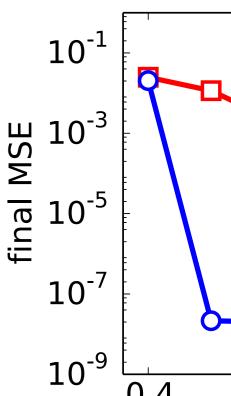


The transition becomes sharper and sharper

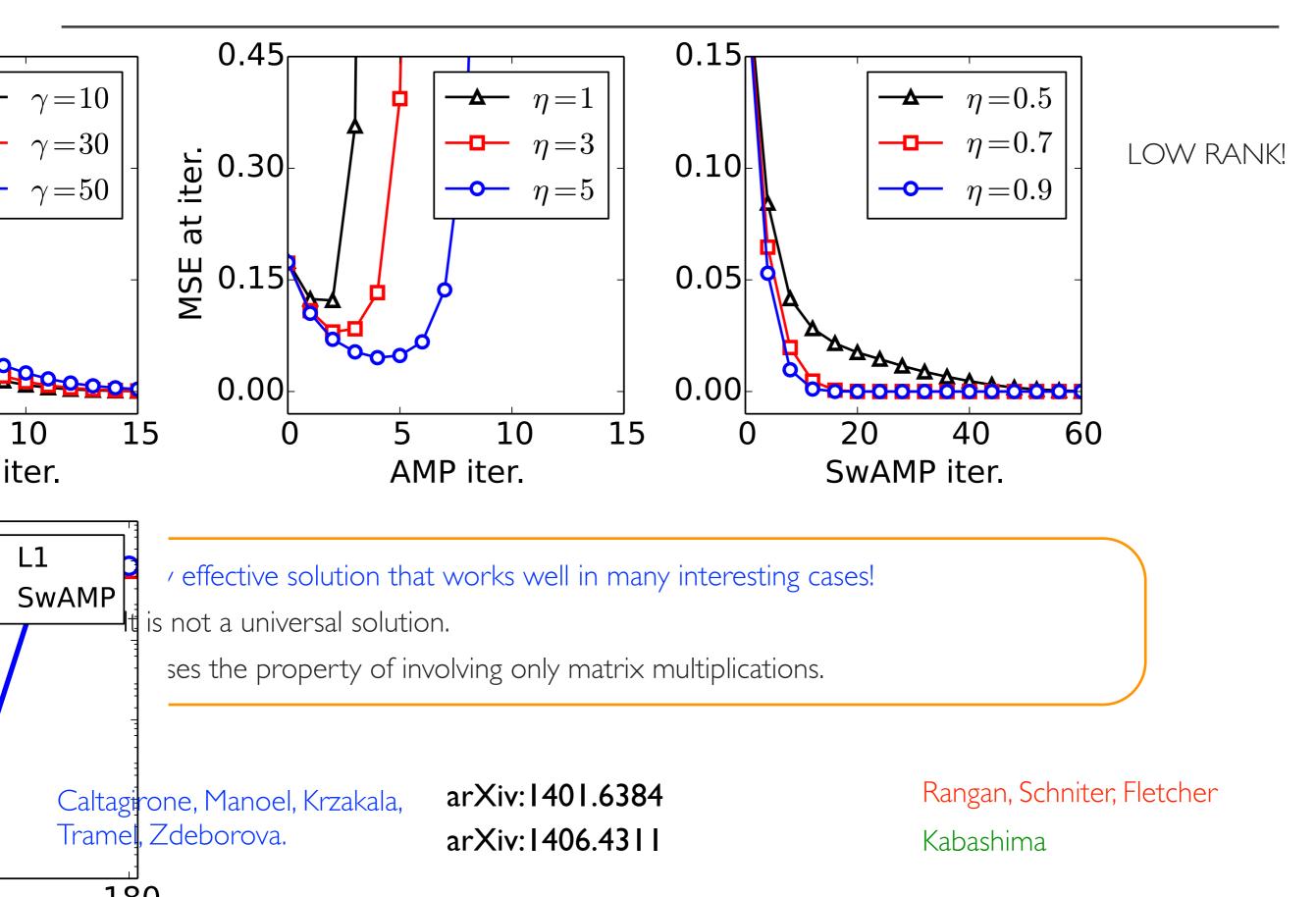
It is expected to move towards the second critical value and behave similarly to the density evolution.



With random sequential update convergence problems disappear.



SwAMP algorithm, a possible solution



- We found that the origin of the convergence problems is an instability of the Nishimori Line
- We provided a possible solution with the SwAMP algorithm.

- Relate this kind of instability in the density evolution to the shape of the replica potential.
- Perform the same kind of analysis for the case of dictionary learning.

THANKYOU!