# q-gram analysis and urn models

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# Approximate pattern matching and the Jokinen-Ukkonen lemma

Def: q-gram any word of fixed size q

#### Edit operations over strings

- substitution  $(l_1 \rightarrow l_2)$   $aabdd \rightarrow aadcc$
- insertion  $(| \rightarrow l)$   $aa|dd \rightarrow aaecc$
- suppression  $(l \rightarrow |)$   $aaedd \rightarrow aa|cc$

#### Edit distance $\delta(S_1, S_2)$ between two strings $S_1$ and $S_2$

- minimum number of edit operations transforming  $S_1$  into  $S_2$ 

#### Jokinen-Ukkonen 1991 (loose version)

if  $|S_1| = m$  and  $\delta(S_1, S_2) \leq k$ , then at least m + 1 - (k+1)q of the m - q + 1 q-grams of  $S_1$  occur in  $S_2$ 

# Example

$$S_1 = aaabaaab$$

$$S_2 = aaacaaaa$$

$$m = 8, \quad \delta(S_1, S_2) = 2 \rightarrow k = 2$$

$$2 - \operatorname{grams}(S_1) = \{\{aa, aa, ab, ba, aa, aa, ab\}\}$$

$$Q_{S_1,S_2} = 2 - \operatorname{grams}(S_1)$$
 present in  $S_2 = \{\{aa, aa, aa, aa\}\}$ 

Jokinen-Ukkonen

$$|Q_{S_1,S_2}| \ge m+1-(k+1)q$$
  
 $4 \ge 8+1-(2+1)2=3$ 

Beware of the asymmetry:  $|Q_{S_2,S_1}| = 5$ 

# Application

When searching a pattern with errors in a text, slide over the text a window of same size as the pattern and discard windows which do not contain enough q-grams of the pattern

#### Aim of this work

Study of two statistics of q-grams in random sequences:

- number of "repeated" q-grams (number of q-grams occurring at least twice, without counting multiplicities

$$S = aaaabaaaabbb, \quad q = 2$$
 
$$Q_{\mbox{repeated}} = \{aa, ab, bb\} \quad |Q| = 3$$

 number of common q-grams to two sequences, without counting multiplicities

$$S_1 = aaaabaaaabbb$$
 
$$S_2 = aaaacaaaacbb$$
 
$$q = 2 \qquad Q_{\mathrm{common}} = \{aa, bb\} \quad |Q| = 2$$

(Remark: symmetrical counting)

- Jokinen-Ukkonen statistics

Bernoulli non-uniform model for the sequences

# A heuristic approach

#### Dependent model

```
FGSEWWTYURR ... OOUYJREFDKB

FGSEWWTYU ...
GSEWWTYU ...
```

. . .

EWWWTYU ...

SEWWTYU ...

#### Independent model

```
TTG
GSE
UHI
ROY
...
sequence length l = n + q - 1 \Rightarrow n q-grams
```

- 1. analyse the independent model
- 2. perform simulations for the dependent model and compare with the independent model

# Repeated q-grams

#### Equivalent problems

Input: an alphabet  $\Sigma$   $(|\Sigma|=s)$ , an integer q, a random sequence S of size n+q-1

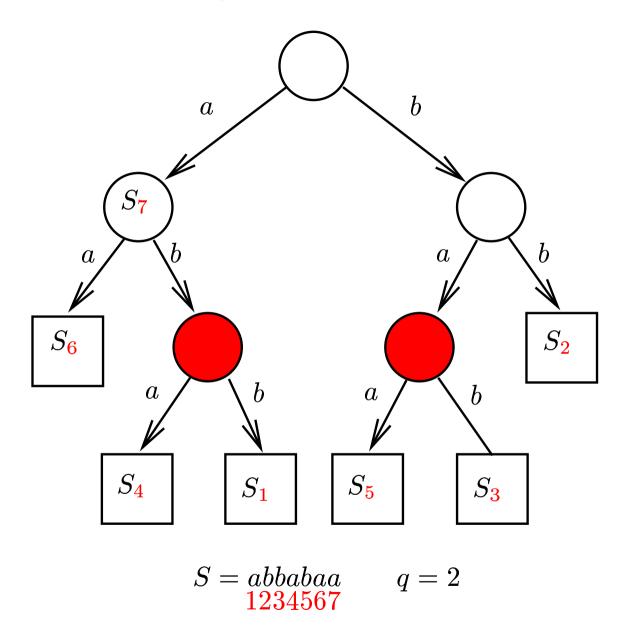
#### Dependent model

- 1. number of repeated q-grams
- 2. number of internal nodes at depth q of the suffix-tree build on S
- 3. number of self-intersections of a random walk of length n over the de Bruijn graph B(s,q)

#### Independent model

- 1. number of repeated q-grams
- 2. number of internal nodes at depth q of a trie build with n random keys over  $\Sigma$
- 3. number of self-intersections of a random walk of length n over a complete graph  $K(s^q)$
- 4. number of urns containing more than one ball in a system of  $s^q$  urns in which n balls are thrown

#### Suffix-trees



 $Q_{\text{repeated}} = \{ab, ba\}$ 

|Q| = 2 = number of internal nodes at depth q

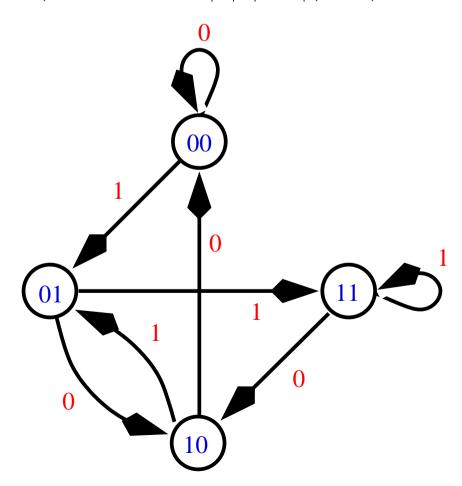
#### DE BRUIJN graphs

#### DE BRUIJN graph B(s,q)

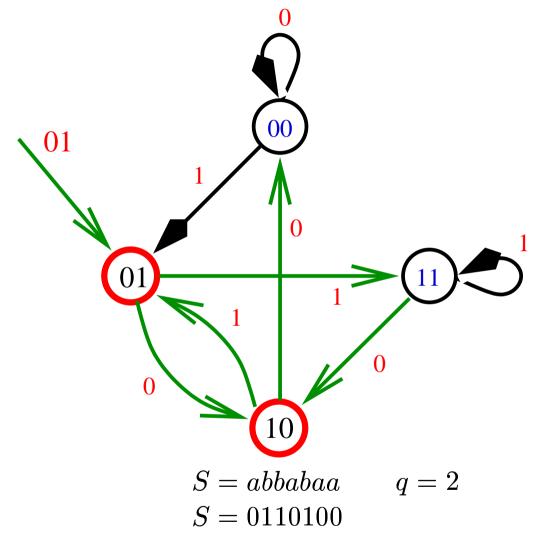
Vertices:  $V = \{0, 1, 2, \dots, s^q - 1\}$ 

Edges:  $E = \{(v_i, v_j)\}$  with

 $v_j = s \times v_i \pmod{s^q} + x, \quad x = 0|1|2|\dots|(s-1)|$ 



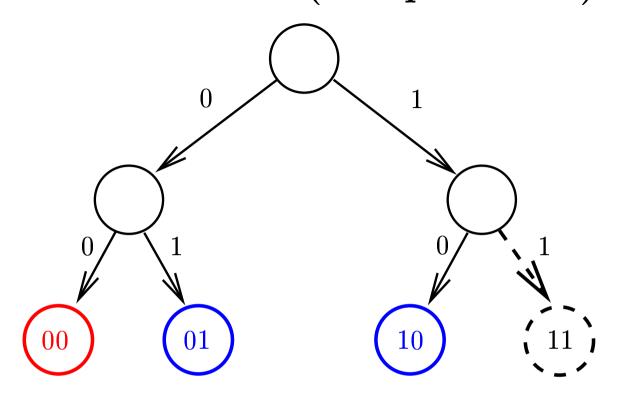
# Random walks over DE BRUIJN graphs



 $Q_{\rm repeated} = \{\textbf{01}, \textbf{10}\}$ 

|Q| = 2 = number of vertices accessed more than once

# Trie and urns (indep. model)



keys = [00, 01, 00, 00, 10, 00]  $Q_{repeated} = \{00\}$ 

|Q|=1= number of nodes at depth q containing more than one key equivalent to a system of 4 urns

 $key \leftrightarrow number of urn$ 

$$\ker = w_{q-1}w_{q-2}\dots w_0$$

$$\Pr(\operatorname{urn}_i) = \Pr(\ker_i) = \prod_{0 \le i \le q-1} \Pr(w_i)$$

#### Previous results

- Guibas and Odlyzko 1981, Rahman and Rivals 2000, 2003
   enumeration of autocorrelations, missing words
- Szpankowski and Jacquet 1994
   asymptotically, the distributions of path lengths of suffix-trees and of tries of same size are equal
  - J. Fayolle 2002, same result, but for the expectation
- Szpankowski and Sutinen 1999
   phase transition in q-gram filtration
- urn models: numerous results
   Johnson and Kotz 1977, Kolchin et al. 1978, Drmota et al. 2001,
   Flajolet et al. 2003

# Analysis of the urn model

 $X_n$  random variable counting the number of urns without collisions when n balls are thrown in the system of  $m = s^q$  urns

 $Y_n = m - X_n$  counts urns with collisions

G.F.

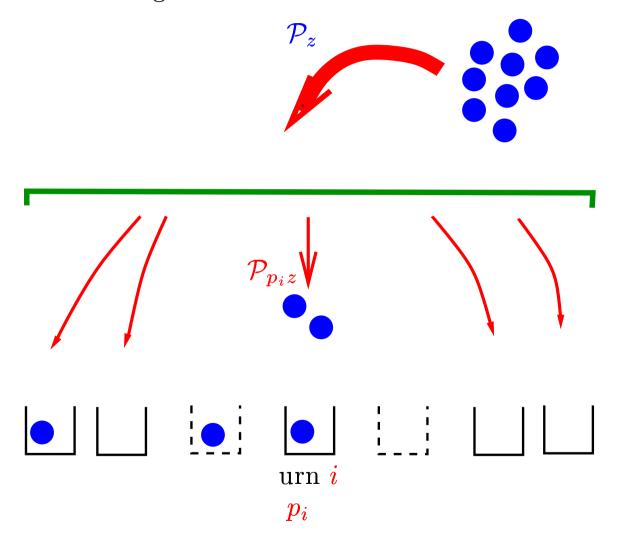
$$F(z, u) = \sum \Pr(X_n = k) u^k \frac{z^n}{n!}$$

differentiations with respect to u

- $\rightarrow$  gen. functions of moments of  $X_n$
- $\rightarrow$  extraction of *nth* Taylor coefficient and asymptotic evaluation

#### Poissonization

do not throw exactly n balls in the urns, but throw a random number of balls following a Poisson distribution.



The urns behave independently of each other

# Poissonization - Depoissonization

 $\mathcal{P}_{p_i z}$  balls in urn i.

 $Pr(\text{no collision}) = e^{-p_i z} (1 + p_i z).$ 

*u* counts the urns without collisions

b.g.f. for urn i under the Poisson model

$$\phi_i(z, u) = e^{-p_i z} ((1 + p_i z) u + e^{p_i z} - 1 - p_i z)$$

for the system of urns (Poisson again)  $\Phi = \prod \phi_i$ 

$$\Phi(z, u) = e^{-z} \prod_{0 \le i \le m-1} \left( e^{p_i z} + (u - 1)(1 + p_i z) \right)$$

"exact" g.f. 
$$F(z, u) = \sum_{n=0}^{\infty} f_n(u) \frac{z^n}{n!}$$

$$\Phi(z,u) = \sum_{n>0} f_n(u) \frac{z^n}{n!} e^{-z} \Leftrightarrow f_n(u) = [z^n] n! e^z \Phi(z,u)$$

$$\Rightarrow F(z, u) = \prod_{0 \le i \le m-1} \left( e^{p_i z} + (u - 1)(1 + p_i z) \right)$$

# Expectation and standard dev.

$$\mu_n = \mathbf{E}(X_n) \qquad m_n^{(2)} = \mathbf{E}(X_n^2)$$

$$m(z) = \sum \mu_n z^n = \left. \frac{\partial F(z, u)}{\partial u} \right|_{u=1}$$

$$m^{(2)}(z) = \sum m_n^{(2)} z^n = \left. \frac{\partial}{\partial u} u \frac{\partial F(z, u)}{\partial u} \right|_{u=1}$$

extract  $[z^n]m(z)$  and  $[z^n]m^{(2)}(z)$  + asymptotics

when  $n \times p_i \to \theta_i$ 

$$\mu_{n} = \sum_{i} \left( e^{-\theta_{i}} (1 + \theta_{i}) + \frac{1}{2n} e^{-\theta_{i}} \theta_{i}^{2} (1 - \theta_{i}) + O\left(\frac{1}{n^{2}}\right) \right) \quad \text{and} \quad \gamma_{n} = m - \mu_{n}$$

$$\sigma_n^2 = m_n^{(2)} - \mu_n^2 \approx \sum_i e^{-\theta_i} (1 + \theta_i) \left( 1 - e^{-\theta_i} (1 + \theta_i) \right) - \frac{1}{n} \left( \sum_i \theta_i^2 e^{-\theta_i} \right)^2$$

# Poisson convergence (Chen-Stein)

number of empty urns: Barbour - Holst 1989

$$I_k = \begin{cases} 1 \text{ if urn } k \text{ empty} \\ 0 \text{ elsewhere} \end{cases} \qquad W = \sum_k I_k \quad \mu = \mathbf{E}(W)$$

$$\text{urn } k$$

$$p_k$$

- (1) empty urn k by throwing the balls into the other urns
- (2) coupling: after this operation

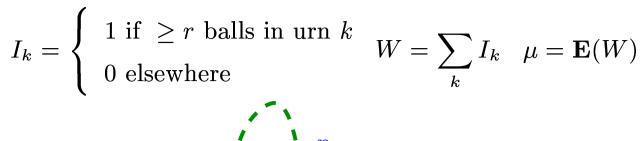
$$\begin{cases}
J_{ik} = \begin{cases}
1 \text{ if urn } i \text{ empty} \\
0 \text{ elsewhere}
\end{cases} \Rightarrow J_{ik} \leq I_{ik} \ (i \neq k)$$

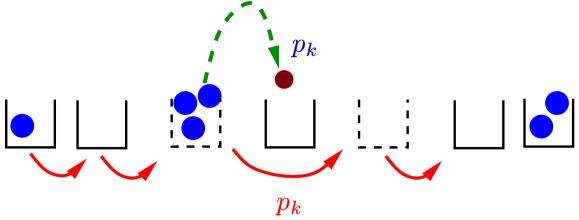
$$I_{ik} = I_i \ \forall k$$

$$\mathcal{L}(J_{1k}, \dots, J_{mk}) = \mathcal{L}(I_{1k}, \dots, I_{mk} | I_k = 1)$$

$$\Rightarrow d(W, \mathcal{P}_{\mu}) \leq \min(1, \mu) \left(1 - \frac{\mathbf{Var}W}{\mu}\right)$$

# Poisson convergence (r-collisions)





if less than r balls in urn k

repeat until there are  $\geq r$  balls in urn kfor all urns  $i \neq k$ for each ball in urn ithrow it into urn k with proba.  $p_k$ 

number of iterations finite with proba. 1 coupling + same proof as Barbour and Holst

$$\Rightarrow d(W, \mathcal{P}_{\mu}) \leq \min(1, \mu) \left(1 - \frac{\mathbf{Var}W}{\mu}\right)$$

# Dependent model

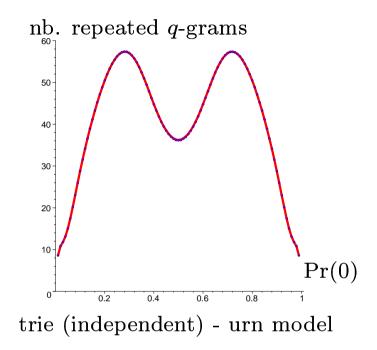
Th: the language of words containing e repeated q-grams is rational, for all e

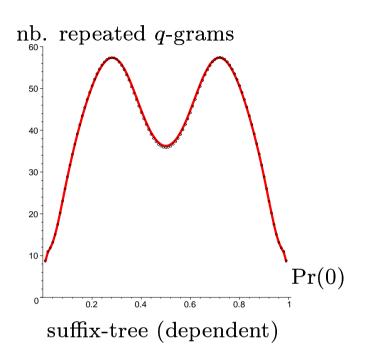
- 1. consider the DE BRUIJN directed graph B(s,q) as an automaton  $(\Sigma,Q,0,\delta,F=Q)$  where the states of Q (vertices) are naturally numbered from 0 to  $s^q-1$  and all states are terminal
- 2. Consider  $3^{s^q}$  copies of B(s,q) corresponding of all combinations of labelling with  $\lambda = 0|1|2$  of the vertices of B(s,q)
- 3. Number the copies along the numbering of the states and the labels:  $B_N(s,q) \Leftrightarrow \text{label of vertex } n \text{ is the } nth \text{ digit of } N \text{ in base } 3.$
- 4. build a (huge) automaton  $(\Sigma, \mathcal{Q}, 0_0, \Delta, \mathcal{Q})$ where  $\mathcal{Q} = \{0, 1, \dots, s^q - 1\} \times \{0, 1, \dots, 3^{s^q} - 1\}$  (notation [n, N]) by connecting the copies

$$\begin{cases} \lambda = 0, 1 : \Delta([n, N_1], l) = [\delta(n, l), N_2] \ (N_2 = N_1 + 3^n) \\ \lambda = 2 : \Delta([n, N_1], l) = [\delta(n, l), N_1] \end{cases}$$

- 5. mark with letter u all transitions changing a label from 1 to 2 (first repetition)
- 6. Chomski-Schützenberger algorithm for marked automata

# Experimental comparisons - Exp





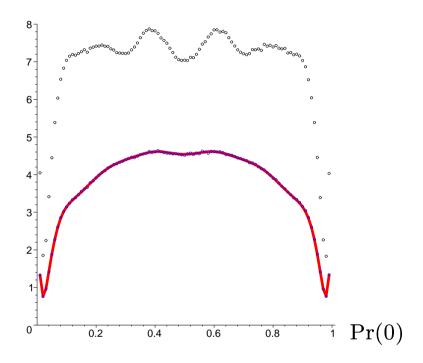
$$n = 300$$
  $\Sigma = \{0, 1\}$   $s = 2$   $q = 10$ 

solid lines: theoretical curve for the trie

dots: simulations

# Experimental comparisons - Std. dev.

repeated q-grams



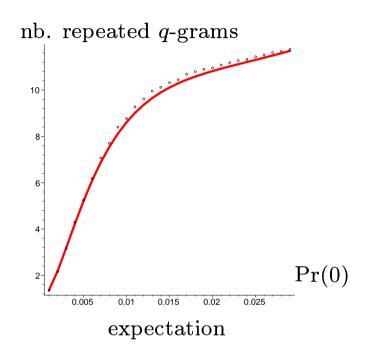
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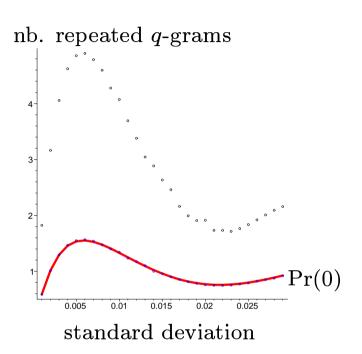
theoretical  $\sigma$  - trie (solid line)

simulations for  $\sigma$  trie (blue circles)

simulations for  $\sigma$  suffix-tree (black circles)

# Small p





$$n = 300$$
  $s = 2$   $q = 10$ 

$$(0.995 + 0.005u)^{300} = 0.2223 + 0.3351u + 0.2518u^2 + 0.1257u^3 + 0.047u^4 \dots$$

# Common q-grams to 2 sequences

#### Equivalent problems

Input: an alphabet  $\Sigma$  ( $|\Sigma| = s$ ), an integer q, 2 random sequence  $S_1$  and  $S_2$  of size n + q - 1

#### Dependent model

- 1. number of repeated q-grams
- 2. number of bicolor nodes at depth q when superposing colored suffix-trees build on  $S_1$  and  $S_2$
- 3. number of intersections of two random walk of length n over the de Bruijn graph B(s,q)

#### Independent model

- 1. number of repeated q-grams
- 2. number of bicolor nodes at depth q when superposing two colored tries build each with n random keys over  $\Sigma$
- 3. number of intersections of two random walks of length n over a complete graph  $K(s^q)$
- 4. number of urns with bicolor collisions in a system of  $s^q$  urns in which n black and n white balls are thrown

#### Previous results

P. Flajolet, P. Kirschenhofer, and R. F. Tichy - 1988, W.
 Szpankowski -1993

asymptotically, all words of size  $\log(n)/H$  are present in a text of size n

(H Renyi-entropy of the alphabet)

 $H = \log \omega_{\min}$  where  $\omega_{\min}$  is the minimum of the probability of the letters of the alphabet

# Analysis of the urn model

#### - g.f. and moments

double poissonization-depoissonization

$$F(z,t,u) = \prod_{0 \le i \le s^q - 1} \left( e^{p_i(z+t)} + (u-1)(e^{p_i z} + e^{p_i t} - 1) \right)$$

z black balls , t white balls u bicolor collisions

$$\mu_{n} = m - [z^{n}t^{n}] \frac{\partial F(z, t, u)}{\partial u} \Big|_{u=1}$$

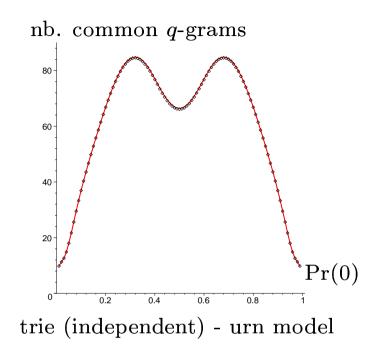
$$= m - \sum_{i} \left( e^{-\theta_{i}} (2 - e^{-\theta_{i}}) - \frac{\theta_{i}^{2} e^{-\theta_{i}}}{n} (1 - e^{-\theta_{i}}) \right) + o(1)$$

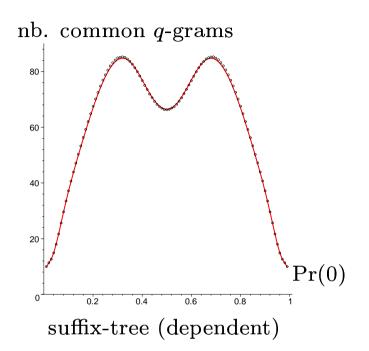
$$\sigma_n^2 \approx \sum_i e^{-\theta_i} (2 - e^{-\theta_i}) \left( 1 - e^{-\theta_i} (2 - e^{-\theta_i}) \right)$$
$$- \frac{2}{n} \left( \left( \sum_i \theta_i e^{-\theta_i} \left( 1 - e^{-\theta_i} \right) \right)^2 - \sum_i \theta_i^2 e^{-2\theta_i} (1 - e^{-\theta_i})^2 \right)$$

#### Poisson convergence

Chen-Stein + coupling (reverse Barbour-Holst)

# Experimental comparisons - Exp





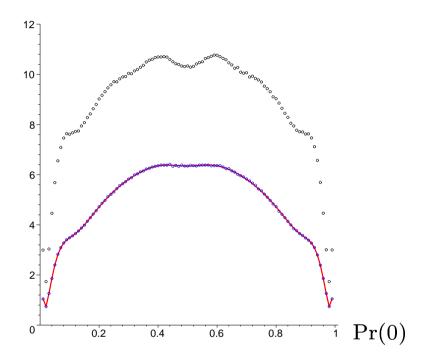
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solid lines: theoretical curve for the trie

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# Experimental comparisons - Std. dev.

common q-grams



$$n=300$$
  $\Sigma=\{0,1\}$   $s=2$   $q=10$  theoretical  $\sigma$  - trie (solid line) simulations for  $\sigma$  trie (blue circles) simulations for  $\sigma$  suffix-tree (black circles)

#### Cost of summations

$$\Sigma = \{1, 2, 3, 4\}, s = |\Sigma| \qquad m = s^q$$

group urns by families of urns with equal probability

|w| = q,  $|w_i| = q_i$  number letters equal to i,

$$q = q_1 + q_2 + q_3 + q_4$$
 population of  $(q_1, q_2, q_3, q_4) = \frac{q!}{q_1! q_2! q_3! q_4!}$ 

#### Number of families $C_{q,s}$ (cost of summation)

 $C_{q,s} = \text{compositions with} s \text{ summands } \geq 0 \text{ of } q$ 

= compositions with summands > 0 of q+s

$$C_s(z) = \left(\frac{z}{1-z}\right)^s$$

$$C_{q,s} = \left[z^{q+s}\right] \left(\frac{z}{1-z}\right)^s = \left(q+s-1\atop s-1\right)$$

ADN:  $C_{10,4} = 286$  Proteins:  $C_{3,20} = 1540$ 

# Computing the moments (repeated q-grams)

The values of  $q_1$  to  $q_{i-1}$  have been computed previously when Procedure Calcsum is entered and d = s - i.

 $s = |\Sigma|$  and q are handled as global constants.

Procedure Calcsum 
$$(f, d, n, \phi)$$
:

$$i = s - d \qquad \quad u = \sum_{k=1}^{i-1} q_k$$

If d > 1 Then

For 
$$j$$
 To  $s-u$  Do

$$q_i = j$$
  $f = \mathbf{Calcsum}(f, d-1, n, \phi)$ 

End of for

Else

$$q_s = q - \sum_{k=1}^{s-1} q_k$$

$$f = f + \frac{q!}{q_1! q_2! \dots q_s!} \phi(\theta_{q_1, \dots, q_s}, n)$$

End of if

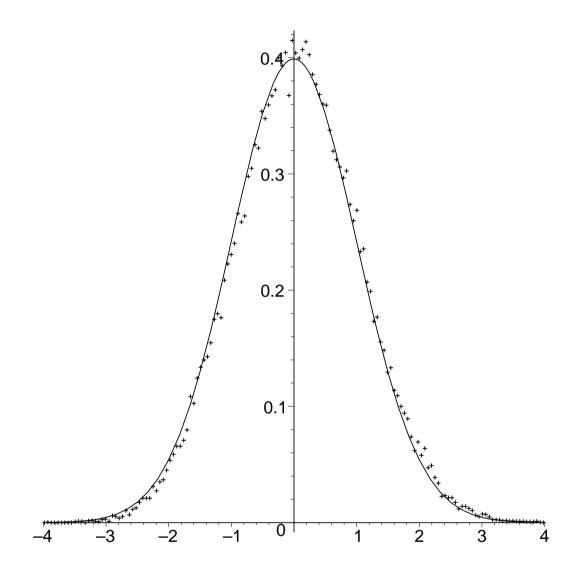
Return (f)

End of procedure

$$\theta_{\xi} = \theta_{q_1, \dots q_s} = n \times \omega_1^{q_1} \omega_2^{q_2} \dots \omega_s^{q_s}$$

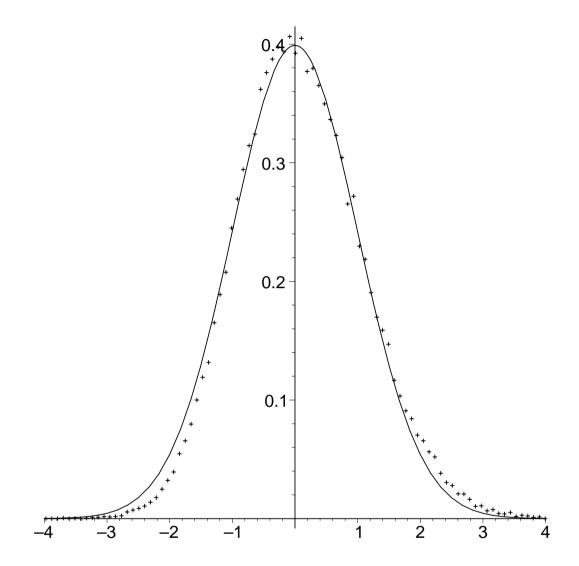
$$\phi_1 = \left( e^{-\theta_{\xi}} (1 + \theta_{\xi}) + \frac{1}{2n} e^{-\theta_{\xi}} \theta_{\xi}^2 (1 - \theta_{\xi}) \right)$$
$$\mu_n = m - \mathbf{Calcsum}(0, s, n, \phi_1)$$

$$\mu_n = m - \mathbf{Calcsum}(0, s, n, \phi_1)$$



 $n_1 = n_2 = 1000, q = 12, \Sigma = \{0, 1\} \ p_0 = p_1 = 0.5$ 50000 simulations

K number of common 12-grams Plot of the normalized variable  $\widehat{K}$  versus  $\mathcal{N}(0,1)$ 



$$n_1 = n_2 = 1000, \ q = 12, \ \Sigma = \{0, 1\} \ p_0 = 0.1 \quad p_1 = 0.9$$
50000 simulations

K number of common 12-grams Plot of the normalized variable  $\widehat{K}$  versus  $\mathcal{N}(0,1)$ 

# Jokinen-Ukkonen statistics (common q-grams)

#### urn model

z counts black balls,  $p_i = \Pr(\text{black ball falls in urn } i)$ t counts white balls,  $x_i = \Pr(\text{white ball falls in urn } i)$ 

double poissonization-depoissonization, g. f. for one urn:  $e^{p_i z} \times e^{x_i t}$ 

*u* counts the total number of black balls that are present in urns containing at least one white ball

$$\begin{bmatrix} 1 & (p_i z) & \dots & \frac{(p_i z)^i}{i!} & \dots \\ (x_i t) & \mathbf{u}(p_i z)(x_i t) & \dots & \frac{\mathbf{u}^i (p_i z)^i}{i!} (x_i t) & \dots \\ \dots & \dots & \dots & \dots & \dots \\ \frac{(x_i t)^j}{j!} & \mathbf{u}(p_i z) \frac{(x_i t)^j}{j!} & \dots & \frac{\mathbf{u}^i (p_i z)^i}{i!} \frac{(x_i t)^j}{j!} & \dots \\ \dots & \dots & \dots & \dots & \dots \end{bmatrix}$$

$$F(z,t,\mathbf{u}) = \prod_{0 \le i \le s^q - 1} e^{p_i \mathbf{u}z + x_i t} - e^{p_i \mathbf{u}z} + e^{p_i z} = \sum f_{\mathbf{k}ab} \mathbf{u}^{\mathbf{k}} z^a t^b$$

 $f_{kab} = \Pr(k \text{ black balls in urns with at least 1 white ball}$  when a white and b black balls are thrown).

# Expectation and Standard Deviation

$$p_i = x_i, \quad a = b = n$$
 $n \to \infty, \quad n \times p_i \to \theta_i$ 

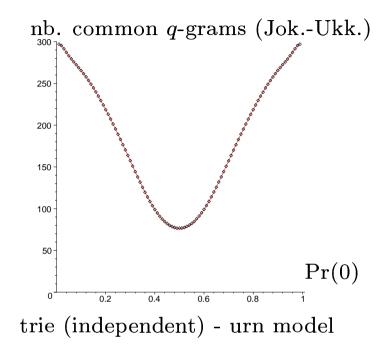
$$\kappa_i = \sum_i \theta_i \left( 1 - e^{-\theta_i} \left( 1 - \frac{\theta_i^2}{2n} \right) \right)$$

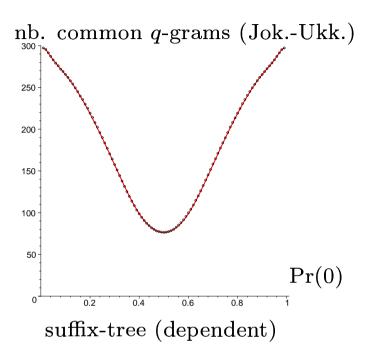
$$\mu_n pprox \sum_i \kappa_i$$

$$\sigma_n^2 \approx \sum_i \kappa_i(\theta_i - \kappa_i)$$

$$-\frac{1}{n} \left( \left( \sum_i \theta_i \left( 1 - e^{-\theta_i} \right) \right)^2 + \left( \sum_i \theta_i^2 e^{-\theta_i} \right) \right)$$

# Experimental comparisons - Exp





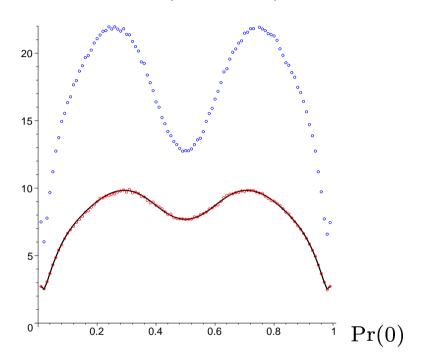
n = 300  $\Sigma = \{0, 1\}$  s = 2 q = 10

solid lines: theoretical curve for the trie

dots: simulations

# Experimental comparisons - Std. dev.

common q-grams (Jok.-Ukk.)



$$n=300$$
  $\Sigma=\{0,1\}$   $s=2$   $q=10$  theoretical  $\sigma$  - trie (solid line) simulations for  $\sigma$  trie (blue circles)

simulations for  $\sigma$  suffix-tree (black circles)